



City Research Online

City St George's, University of London

Citation: Laušev, J., Stojanovic, A. & Todorovic, N. (2011). Determinants of debt rescheduling in Eastern European countries. *Economic Annals*, 56(188), pp. 7-31. doi: 10.2298/eka1188007l

This is the accepted version of the paper.

This version of the publication may differ from the final published version. To cite this item please consult the publisher's version.

Permanent repository link: <https://city-ref-test.eprints-hosting.org/id/eprint/7692/>

Link to published version: <https://doi.org/10.2298/eka1188007l>

Copyright and Reuse: Copyright and Moral Rights remain with the author(s) and/or copyright holders. Copies of full items can be used for personal research or study, educational, or not-for-profit purposes without prior permission or charge, unless otherwise indicated, provided that the authors, title and full bibliographic details are credited, a hyperlink and/or URL is given for the original metadata page and the content is not changed in any way. For full details of reuse please refer to [City Research Online policy](#).

DETERMINANTS OF DEBT RESCHEDULING IN EASTERN EUROPEAN COUNTRIES

Jelena Lausev¹, Aleksandar Stojanovic² and Natasa Todorovic³

ABSTRACT

This study utilizes Panel Logit Models applied to a set of macroeconomic, financial and political variables to estimate debt rescheduling probabilities of 15 Eastern European countries during the transition period from 1990-2005. These transition economies became a very attractive region for foreign investments. Specifically, the region became the largest recipient of net non-FDI flows among all emerging market regions in 2005. Therefore, it is relevant for policy makers, institutional and private foreign investors to investigate factors that influence debt rescheduling probabilities, as these may directly affect the size of and the return on the investments in these countries. Our findings suggest that the policy efforts focused at reducing government expenditures, attracting foreign direct investments, increasing export revenues and keeping a good repayment record result in low debt rescheduling probabilities and, in turn, decrease the cost of debt for these countries. This is the common finding for all countries in the sample, including those that became EU members.

Key words: Country Debt, Rescheduling, Eastern Europe, Transition, Panel Logit Model

JEL Classification : F33, F34

¹ Faculty of Economics, University of Belgrade, Kamenicka 6, 11000 Belgrade, Serbia

² University of Greenwich Business School, Old Royal Naval College, Park Row, London SE10 9LS, UK

³ **Corresponding author:** Cass Business School, 106 Bunhill Row, London EC1Y 8TZ, UK;
email: n.todorovic@city.ac.uk

1. INTRODUCTION

All major international investors and banks use sovereign debt default rates to price sovereign bonds and loans as well as to determine a country risk. Furthermore, in recent years many developing countries have negotiated new loan repayment schedules with their government and commercial bank lenders. The aim of this study is to specify a model which would allow prediction of sovereign default/rescheduling rates with a higher accuracy, particularly tailored for the investors interested in the Eastern Europe (EE) region.

The group of 15 EE countries used in this study share similarities in respect of the social, political, geographical, economic and cultural characteristics. Most of the recent studies of sovereign debt problems looked at the middle income countries of Latin America, Asia or Africa, but none of them paid special attention to emerging markets of Eastern Europe, and the period of their transition to market economies. Several earlier studies include data for some of these countries in their multi country data set, however, due to specific common characteristics of EE countries, it is questionable whether conclusions from previous studies looking at all emerging markets can be directly applied to EE countries. Therefore, this paper will focus on EE countries as the very first attempt to model debt rescheduling probabilities for this region.

This paper aims to answer the following questions: (1) What are the most important determinants of EE sovereign debt rescheduling? (2) How accurately those determinants predict sovereign debt rescheduling? (3) How to specify a model that would allow prediction of EE sovereign rescheduling with a higher accuracy? Answers to these questions could be used to derive implications for policy makers and potential creditors and investors in the region.

The countries included in the study are: Albania, Bosnia and Herzegovina, Bulgaria, Croatia, Czech Republic, Estonia, Hungary, Latvia, Lithuania, Macedonia, Poland, Romania, Russian Federation, Serbia, Montenegro and Slovak Republic. The regions' debt rescheduling probabilities are tested for the effects of wide range of macroeconomic, financial and political indicators.

2. LITERATURE REVIEW

Series of financial crises and defaults starting from the late 1970s and early 1980s period (see e.g. Camdessus, 1986 and Cuddington and Smith, 1985), and later incidents culminating in Argentina's default in 2001, prompted an intense debate about how to prevent crises. As pointed out by Manasse, Roubini and Schimmelpfennig (2003) the literature on debt crisis falls into four broad categories: theoretical models of sovereign default, empirical studies of the determinants of debt crisis, empirical studies of the predictive power of credit ratings and empirical studies of the determination of spreads. Given the aims of our study, the literature we review is focused on the determinants of debt rescheduling and the models used to predict it.

One of the earlier studies by Frank and Cline (1971) applied discriminate analysis to show that it is possible to obtain a very high prediction rate using only two factors, the debt service ratio and the average maturity of debt. Among the first studies using logit analysis were Feder and Just (1976) followed by Feder, Just and Ross (1981) and Feder and Uy (1984). The authors point out that the analysis of the growth record of many economies indicates that foreign capital is an important factor in the process of economic development. Eaton and Gersovitz (1981) developed a model of borrowing with default, showing that lenders establish debt ceilings above which they are unwilling to increase loans. Further, McFaden et al. (1985) estimate crisis probabilities using data on 93 countries from 1970-1982, concluding that although an econometric model can help to explain sovereign debt problems, it is extremely difficult to develop an early warning system that would predict debt crises in advance as countries may experience heterogeneous repayment problems. They find that the debt burden, the level of per capita income, real GDP growth, and liquidity measures such as non-gold reserves are significant predictors of debt crisis, while the changes in real exchange rates are not.

In most of the empirical studies two groups of main variables were constantly included in econometric analyses: short-term or liquidity factors (also known as traditional debt or financial ratios: debt/GDP, reserves/imports and debt service/export) and long term variables (GDP growth, investments, inflation). For example, Manasse, Roubini and Schimmelpfennig (2003) find solvency problems measured through high total external debt to GDP ratio and liquidity problems measured through short term debt to be highly associated with the debt

rescheduling. Hemming and Petrie (2000) present a discussion on fiscal sustainability and develop vulnerability indicators. Hemming and Chalk (2000) for the first time systematically examine the link between fiscal and external sustainability. Furthermore, Detragiache and Spilimbergo (2001) study the importance of liquidity factors such as short-term debt, debt service, and the level of international reserves in predicting debt crises, and find that less liquid countries are more likely to default. Additionally, they find that countries that are more open are in a better position to service debt and that overvalued exchange rate hurts the future export performance of the country.

Krugman (1987) built theoretical model that highlights the importance of “debt overhang” of the developing nations and derive policy implications about the desirability of forgiving or refinancing such debt. Berg and Sachs (1988) found that higher income inequality is a significant predictor of a higher probability of debt rescheduling in the middle income countries. Also, the open trade regime is a significant predictor of a reduced probability of debt rescheduling. As they pointed out, earlier studies, such as McFadden et al. (1985), identify variables that are more the symptoms of the crises rather than their fundamental causes.

Political risk and institutional variables are also found to play a key role in determining debt rescheduling among emerging markets countries. Balkan (1992) was the first to find an inverse relationship between rescheduling probabilities for one country and its level of democracy, as well as a direct relationship between probability of rescheduling and political instability. More recent evidence of the significance of political variable can be found in Haque, Nelson and Mathieson (1998). Brewer and Rivoli (1997) showed that the inclusion of political variables into forecasting models improves the correct prediction rate of rescheduling probability by nine to twelve percent for the earlier periods and 18 to 35 percent for later periods. Citron and Nickelsburg (1986) point out that the political variable is in fact a non diversifiable risk and identify waves of default which tend to occur in approximately thirty year cycles and typically involve many countries. Georgievska et al. (2009) confirm the importance of political factors in determining debt rescheduling probabilities in emerging markets. Kraay and Nehru (2004) were the first to point that quality of institutions and policies has a key impact on rescheduling. Further, Butler and Fauver (2006) found that the quality of a country’s legal and political institutions play a vital role in determining sovereign

credit ratings. Reinhart and Rogoff (2004) find that the flows of capital from rich countries to poor countries are largely governed by sovereign default probabilities.

A number of studies is also assessing the impact of past debt repayment record on the future rescheduling. For example, Carmen (1992), McFadden et al (1985) and Aylward and Thorne (1998) find that the poor debt repayment history is a strong indicator of future problems.

Ciarlone and Trebeschi (2005) use financial market data to analyze the behavior of CDSs spreads. Schwartz and Zurita (1992) develop a model which determines the optimal amount of debt to borrow based on the production possibilities of the country, time preferences and the risk free interest rate. The study also gives the optimal level of debt default penalty: the higher the penalty that can be imposed on the country in the event of default the lower the probability of default but higher the problem of underinvestment. These arguments are relevant for considering possible new debt authorization to the countries of Eastern Europe.

2.1. EE Countries' Recent Economic Position

The European Central Bank considers investing in EE countries during the transition period as potentially problematic due to two main reasons⁴: 1) credit growth can affect financial stability and financing 'bad' projects will eventually turn into bad debt or result in a build-up of large external debt, and 2) policymakers and market participants need to assess and monitor credit growth developments as in the past they have been associated with the emergence of financial and currency crises.

In the early years of the transition, most EE countries experienced a significant slump in GDP, which was followed by a period of rapid economic growth. In some cases, rapid privatization of the banking sector and expansion of underdeveloped financial markets resulted in lending booms followed by credit crunches and crises that spilled over to the region as a whole. Hoti (2005), for example, analyses country-specific and regional 1) economic, 2) financial, 3) political and 4) composite risk factors for six Balkan countries and documents significant country spill-over effects.

⁴ Financial Stability Review, ECB 2006

Eastern Europe is the only emerging market region with systematic current account deficits in 2005-2006. According to Roubini and Menegatti (2006) countries that have large current account deficits above four per cent of GDP are: Lithuania, Hungary, and Slovakia while Schafer (2006) reports that Estonia and Latvia have current account deficits even above ten per cent which makes these countries similar to Romania, Bulgaria, Croatia, Bosnia and Herzegovina and region of Serbia and Montenegro. At the same time some of these countries are recording high lending growth since most of them finance current account deficits by FDI. Most of the other emerging markets in Asia and Latin America are running current account surpluses. Since many previous episodes of financial crisis in emerging market economies have been associated, among other factors, with large current account deficits it is important to examine EE group of countries separately.

The average growth rate of credit in 2001-2005 period has been very high in all countries in our sample but Poland, Czech Republic and Slovakia (see Egert et al. (2006), Kiss and Nagy (2006), Coricelli et al. (2006), European Commission (2006), ECB Financial Stability Review (2006) and World Bank Regular Economic Report (2007)). In addition, Hungary observed asset bubbles in equity and housing markets. Schafer (2006) indicates that these countries' economic boom is financed by foreign lending and concludes that it may lead to financial crisis.

World Economic Outlook (2006) reports that the region's generally large current account deficits partly reflects favorable investment opportunities in EE region. In some countries capital inflows were associated not just with private sector financial imbalances but also with substantial fiscal imbalances (e.g. Hungary, Poland and Slovakia). However, WEO (2006) stresses that the large net capital inflows are increasing not in the form of FDIs but rather in the form of other (more volatile) flows, including short-term debt. Specifically, the EE region became the largest recipient of net non-FDI flows among all emerging market regions in 2005. This has implications for policymakers and potential investors who need to examine the trade off between the sovereign capacities and the risks involved.

Roubini and Menegatti (2006) distinguish between solvency and liquidity variables in EE region, which are also potential determinants of sovereign debt rescheduling. This study reports high and rising levels of external debt as a share of GDP, especially for Hungary, Croatia, Slovenia, Latvia, Estonia and debt to export ratio for Estonia, Hungary, Latvia,

Poland, Slovenia, Croatia, and Romania. Liquidity risk measured as a share of money in foreign exchange reserves (M2/FX ratio) is well above one for all countries in the sample of this study. In addition, short term debt/FX ratio is rising in Hungary, Croatia, Lithuania, Latvia and Estonia. Gross external financing needs (measured as short term debt plus current account deficit as a share of foreign exchange reserves) are greater than one also for all countries. Public sector vulnerabilities are presented through share of foreign currency public debt in total debt which is greater than 50% in Estonia, Latvia, Lithuania, Croatia, Bulgaria, Romania and Bosnia-Herzegovina. The share of short term public debt in total debt is high (above 30%) in Czech Republic and Poland and medium (20-30%) in Hungary, Croatia, Lithuania and Slovakia. Public debt to GDP ratio is high (above 50%) only in Hungary and Bosnia-Herzegovina. However, more relevant measure of solvency, public debt to government revenues ratio is relatively high (above 100%) in a larger set of countries: Hungary, Poland, Slovakia, Croatia and Bosnia-Herzegovina. Overall, general government deficits are medium to high in Hungary, Slovakia, Croatia, Czech Republic, Lithuania, Slovenia and Romania.

Note that nine of the selected countries in our sample are now members of the European Union (EU): Czech Republic, Estonia, Hungary, Latvia, Lithuania, Poland, Slovak Republic and very recently Romania and Bulgaria. This will be taken into account in our models.

3. DATA SAMPLE

Annual data on rescheduling event and determinants of rescheduling is collected for 15 EE countries over the period 1990-2005. Our study includes countries of Central EE (Hungary, Czech Republic, Slovak Republic, Slovenia and Poland), South EE (Croatia, Albania, Bosnia and Herzegovina, Serbia, Montenegro, Macedonia, Bulgaria and Romania), Baltic countries (Estonia, Latvia, Lithuania) and Russia. The reason for this lies in their common historical heritage and initial economic conditions on which they started period of transition in the early 1990s.

3.1. Defining the Dependent Variable: default or rescheduling event

Previous literature shows that there is no single empirical definition of what should constitute a sovereign debt crisis, default or a rescheduling event (see for example Manasse, Roubini and Schimmelpfennig (2003), Ciarlone and Trebeschi (2005), Eaton and Gersovitz (1981),

Hajivassiliou (1987), Odedokun (1995) and Kraay and Nehru (2004)). Broadly speaking, a default is any failure by a debtor to meet its contractual obligations. The main event of default is missing a scheduled payment of principal or interest.

In this paper the dependent variable is defined as the ‘total amount of debt rescheduled’ as is reported in the World Bank - Global Development Finance 2006 (GDF)⁵. If there was report of rescheduling on any of the component items, we have considered that the country has rescheduled.

In the model that we intend to use, the event of whether or not rescheduling took place is dichotomous and qualitative variable. Therefore, the dependent variable is a binary variable possessing only two possible values or outcomes:

$$\text{Rescheduling}_{it} = \begin{cases} 1 & \text{if sovereign } i \text{ reschedules its debt in year } t \\ 0 & \text{if sovereign } i \text{ does not reschedule its debt in year } t \end{cases} \quad (1)$$

Table 1 summarizes all the countries included in analysis, the period examined for each country and the actual rescheduling observations during the sample periods. The final sample is comprised of 176 observations and 35 rescheduling events.

- Insert Table 1 –

3.2. Variables used as possible determinants of debt rescheduling

We start our analysis with 40 macroeconomic, financial and political indicators that can potentially explain countries' sovereign debt rescheduling probabilities. This large number of variables will then be reduced to the significantly smaller number through several econometric steps, discussed in Section 3. The variables are divided in nine categories, specifically: 1) solvency, 2) liquidity, 3) variables used in currency crisis models, 4) macroeconomic variables, 5) external trade ratios 6) public debt ratios, 7) financial variables,

⁵ including debt stock rescheduled, interest rescheduled capitalized, interest rescheduled official, interest rescheduled private, principal rescheduled official, principal rescheduled private and principal rescheduled. Debt from official creditors includes loans from international organizations (such as World Bank, EBRD or other multilateral and intergovernmental agencies) and loans from governments (bilateral loans). Debt from private creditors include bonds, commercial bank loans from private banks and other private financial institutions and other private credits (from manufacturers, exporters, suppliers of goods and bank credits covered by a guarantee of an export credit agency)

8) past rescheduling record and 9) the political variable proxied by ICRG Index⁶. The list of all variables and the impact that each one is *expected* to have on the probability of rescheduling is shown in Table 2.

- Insert Table 2-

The descriptive statistics associated with these variables for countries that have rescheduled during the sample period, as well as for those that have not, are presented in Table A1 in Appendix 1.

Three main sources of data used to compile the set of variables described in Sections 3.1. and 3.2. are: IMF's International Financial Statistics, World Bank's Global Development Finance and World Bank Development Indicators. The political variable, ICRG index, is provided by PRS Group. The frequency of all data is annual. To ensure predictive power of our models, all potential explanatory variables from Section 3.2. are lagged by one year.

4. EMPIRICAL MODEL SPECIFICATION

Using all 40 variables at the same time, would result in an overly complicated model, with many potentially significant coefficients appearing as statistically indistinguishable from zero, i.e. multicollinearity and over-fitting of the model may cause biased values of the significance test. Therefore, to eliminate less relevant variables, we proceed along a three-stage strategy.

4.1. Stage 1: Selecting Relevant Variables - Univariate Logit

At the first stage, we run univariate panel logit regressions for each of the 40 variables independently from one another. This will enable us to exclude from the final model all the variables that turn out to be insignificant in determining whether the sovereign reschedules its debt or not.

To define a panel logit model, consider a country i observed over T periods of time, where $t = 1, \dots, T$ and $i = 1, \dots, N$. For this country there exists an unobservable⁷ random variable y^*_{it}

⁶ In ICRG index, the political risk rating contributes 50% of the composite rating, while the financial and economic risk ratings each contribute 25%. Each component is assigned a maximum numerical value (risk points), with the highest number of points indicating the lowest potential risk for that component and vice versa.

indicating latent propensity. Whilst we do not observe y^*_{it} directly, we do observe a binary outcome y_{it} such that $y_{it} = 1(y^*_{it} > 0)$ where y_{it} is termed the indicator function taking the value 1 if the condition within parentheses is satisfied, and 0 otherwise. An example of this might be the amount of total debt rescheduled (y_{it}), and the capability of the country to pay its debts (y^*_{it}). y^*_{it} is a function of explanatory variable(s) x_{it} , constant unobserved individual county effects α_i and random error term u_{it} . The following equation represents the above:

$$y^*_{it} = \alpha_i + x'_{it} \beta + u_{it} \quad (2)$$

y_{it} is a dummy variable defined by:

$$y_{it} = \begin{cases} 1 & \text{if } y^*_{it} > 0 \\ 0 & \text{otherwise} \end{cases}$$

For any β , the probability of observing the outcomes y_{it} is conditional on the values of variable x_{it} . For the panel logit model, the probability that a sovereign i will reschedule its debt at time t is given as:

$$\text{Prob}(\text{Rescheduling}_{it} = 1) = \frac{\exp(\alpha_i + x_{it}\beta)}{1 + \exp(\alpha_i + x_{it}\beta)} \quad (3)$$

Through this procedure, we extracted 17 variables, namely: Lagged Rescheduling Event, ICRG index, Export/GDP, Current Account/GDP, Total Debt/GDP, GDP growth rate, Foreign Direct Investment/GDP, General Government Expenditures/GDP, Credits to Private Sectors/GDP, Short Term Debt/Total Debt, Inflation Rate, Exchange Rate Devaluation, Interest Arrears/Total Debt, Principal Arrears/Total Debt, M2/Reserves, Imports/GDP and International Reserves/GDP. However, using 17 variables in the same model can still result in biased coefficient and in turn biased interpretation of which variables increase the probability of debt reheduling.

4.2. Stage 2: Selecting Relevant Variables - Principal Component Analysis

In order to eliminate any strong and persistent multicollinearity between 17 explanatory variables obtained from the univariate panel logit approach, in this next stage we resort to the Principal Component Analysis (PCA). The PCA provides a method for simplification,

⁷ "latent" variable

combining many correlated variables into a smaller number of underlying factors (Hamilton (2004)). More specifically, PCA computes a new set of orthogonal values from a linear combination of the explanatory variables. PCA can be defined as:

$$L_i = \alpha_1 X_1 + \dots + \alpha_n X_n \quad (4)$$

where $\sum_{n=1}^n \alpha = 1$; the variance of the equation is maximized and L_i the explanatory variable is orthogonal.

In order to select how many factors to use, we consider eigenvalues from the PCA. For the full set of 17 explanatory variables, it was found that 5 components (factors) have eigenvalues greater than 1, accounting for 71.24% of the total variability in the data and are therefore regarded as significant for our further analysis. After rotating the factor matrix, from the identification of the highest factor loadings for each variable on each factor, we have determined the dimension of each of the 5 factors (components) named after the variables that dominate in each group:

- Factor 1 – *Solvency (Trade) Dimension*: Total Debt/GDP, Foreign Direct Investment/GDP, Credits to Private Sectors/GDP, Imports/GDP and International Reserves/GDP
- Factor 2 – *Liquidity Dimension*: Interest Arrears/Total Debt, Principal Arrears/Total Debt and M2/Reserves
- Factor 3 – *Currency Crises/Macroeconomic Dimension*: Exchange Rate Devaluation Inflation Rate and GDP growth rate
- Factor 4 – *Public Debt Dimension*: General Government Expenditures/GDP and Short Term Debt/Total Debt
- Factor 5 – *External Debt Dimension*: Current Account/GDP.

Thus, from each dimension (component), we have selected no more than two to three variables (12 in total), according to the highest factor loading, which are sufficient in explaining the whole dimension. The results of rotated factor matrix are presented in Table A2 in Appendix 2.

4.3. Stage 3: Building the Multivariate Panel Logit Model: Stepwise Gradual Method

In this final stage, we apply the Forward Stepwise Gradual Model Based Procedure for the multi-variate panel logit model construction with an aim to develop parsimonious models, i.e. models that would have a high degree of correct predictions and low Type I and Type II errors⁸. The reason for this gradual model building is to control for the omitted variables (see Roubini, Mannasse and Schimmelpfennig (2003) and Ciarlone and Trebeschi (2005) for stepwise model building procedure).

Our Forward Stepwise multivariate panel logit regressions indicate that there is persistence in significance of three variables, namely: Total debt/GDP, Exports/GDP and Current Account/GDP. We refer to those variables as ‘fixed’ variables in further text, as proposed in Sala-i-Martin (1997). ‘Fixed’ variables are used as a basis for the models we examine, while additional variables to be included in those multivariate logit models are selected using forward stepwise procedure. Therefore, we generate two models that we believe are more accurate than others:

- 1) Model I includes: the ‘Fixed’ variables, Foreign direct investments, General Government expenditures, Credits to private sector, Short-term debt and the Lagged rescheduling record.
- 2) Model II includes: the ‘Fixed’ variables, the Lagged rescheduling record, GDP growth, M2/Reserves and ICRG composite index. However, eight of the 15 countries analysed joined the EU during the sample period, which may lead to the improvement of the debt repayment record of a country. Thus, to test whether EU membership influences the probability of debt rescheduling, we add an EU dummy variable to Model II, taking value of 1 when a country becomes a member and zero otherwise, creating Model IIa.

Note that the probability of rescheduling in a multivariate panel logit framework is obtained in the similar manner as in equation (3):

⁸ Type I error occurs when actual defaults/reschedulings are classified by the model as non-defaults/reschedulings. Type II error occurs when actual non-defaults/reschedulings are classified by the model as defaults/reschedulings. The cut-off point for classifying a probability as rescheduling or non-rescheduling is determined by default: cut off point of 0.5

$$\text{Prob}(\text{Rescheduling}_{it} = 1) = \frac{\exp(\alpha_i + \mathbf{x}'_{it}\beta)}{1 + \exp(\alpha_i + \mathbf{x}'_{it}\beta)} \quad (5)$$

Where, in the multivariate panel logit model framework, β is a $(k \times 1)$ vector of parameters associated with the transposed vector of variables $\mathbf{x}'_{it}$ selected for the model. Finally, following Hausman test results, we apply random-effect logit model estimation⁹, indicating that state dependence appears to be very important in the case of sovereign debt rescheduling (as found in Aylward and Thorne (1998), Hajivassiliou (1989, 1994), McFadden et al. (1985), Brewer and Rivoli (1997), Carmen (1992) among many others).

5. RESULTS

The results for Model I, Model II and Model IIa are presented in detail in Table 3.

- Insert Table 3 -

All tested variables in Model I are highly statistically significant at 1% and 5% level of significance except for the Lagged dependent variable (rescheduling event) that has a p value slightly above the 10% significance level and Foreign direct investments that are significant at 10%. All significant variables in Model I have the expected sign. Model I and the explanatory variables are jointly statistically significant at 5% level as indicated by p -values of Wald Chi-squared test.

Since the logit model is non-linear, the coefficients obtained by the Model I do not explain the contribution of each variable to the probability of rescheduling, but instead they represent odds ratios which are not easily interpreted. Therefore, we calculate the real marginal effects that each of the variable has on the dependent variable.

A closer look at the marginal effects for Model I shows significant impact of explanatory variables on the probability of debt rescheduling, both 'fixed' explanatory variables and those less typically found in literature on emerging market debt rescheduling probabilities. For example, an increase in one unit of the 'fixed' variable Total Debt/GDP causes an increase in debt rescheduling probability by 2.55%, while a unit increase of Exports/GDP ratio causes decrease in sovereign debt rescheduling probability by 7.03%. Increase in a unit of Current

⁹ **Random effects Panel Logit Model (REM)** is appropriate when there are cross sectional differences or heterogeneity and this heterogeneity is assumed to be not correlated with the regressors of the model.

Account/GDP ratio causes increase in sovereign debt rescheduling probability by 5.06%. This may seem counter intuitive, since it implies that the higher the current account surplus, the more likely is that the country will reschedule its debt. However, for transitional economies, such as the countries in our sample, current account deficits are associated with higher growth and financial integration in the region (see Abiad et. al., 2007), which in turn can lead to improvement of debt repayment record of a country. Foreign Direct Investment/GDP ratio variable shows the level of openness of the country and is a significant indicator of business climate in the country that wants to attract potential investments, so it is not a surprise that its unit increase causes decrease in sovereign debt rescheduling probability by 9.4%. Additionally, countries in transition are well known for their high levels of government spending, therefore, in our sample, a unit increase of Government Expenditure/GDP ratio causes a massive increase in sovereign debt rescheduling probability of 19.25%.

Furthermore, in Table 2, we show that impact of Credits to private sector/GDP and Short term debt share in total debt variables on probability of debt rescheduling can be somewhat ambiguous. Both of those variables are significant at the 5% level and negatively related to probability of debt rescheduling. Specifically, it seems that a country is less likely to reschedule if there is a higher percentage of Short term debt in Total debt. In countries in transition, given their economic and other uncertainties and risks¹⁰, it is not surprising that short-term debt would be perceived as less likely to default than the long-term one. In terms of Credits to private sector/GDP, it is possible that private sector investments in transition economies are perceived to show higher levels of credibility than the public sector. Further, this variable can be an indicator of ownership restructuring and potentials for higher future GDP growth, which in turn decreases the probability of sovereign debt rescheduling. Thus, increase in Credits to private sector can decrease the risk of sovereign debt rescheduling probability as a strong indicator of private sector development in transition countries that investors take into account when they calculate net present values of their investments.

Model II and its explanatory variables are jointly significant at 1% level. Although majority of the variables used in the model are significant, such as the 'fixed' variables (at 1% or 5% level), GDP growth (at 1% level) and lagged rescheduling event (at 1% level), our findings

¹⁰ For further details see Detragiache and Spilimbergo (2001)

show that political factors (ICRG index) and liquidity variable M2/Reserves do not significantly influence debt rescheduling in the region.

Marginal effects in Model II show that one additional debt rescheduling event increases the probability that the country will reschedule in the future by 3.85%. A unit increase in Total debt/GDP and Current Account/GDP increase the probability of debt rescheduling by 3.79% and 9.29% respectively, whereas a unit increase in Export/GDP and GDP growth rate decrease the probability of debt rescheduling by 10.29% and 5.54% respectively. Persistently high influence of Export/GDP variable in both Model I and Model II can be explained by the fact that countries that we are focusing on are mainly small open economies, highly influenced by terms of international trade and finances. At the same time, exports are generating revenues, GDP growth and represent the drivers of stability of these countries. In conclusion, besides the current account and the level of total debt burden, these small open economies' investment opportunities are highly dependent on their past rescheduling records.

One of the results of Model II that we believe requires more attention, is the insignificance of political variable, which was found to be important in many previous studies. One reason for this could be the decrease of potential political risks and stabilisation in the region. Specifically, in the later part of our sample period, most of the countries entered the final phase of transition while some of them joined the European Union. To check for the ultimate importance of being a member of the European Union (perceived in this context as a political factor variable), we extend Model II by including an EU dummy variable to create Model IIa. The insignificant EU dummy coefficient in Model IIa implies that the EU membership does not affect sovereign debt rescheduling probabilities in EE countries during the sample period.

Overall, in terms of determinants of debt rescheduling, it turns out that variables which are more specific for countries in transition have greater impact on debt rescheduling than variables traditionally considered as the most influential ones according to the literature (e.g. political factors, reserves to imports, total debt service payment to exports, inflation, indicators of exchange rate overvaluation, etc.)

In the search for the most accurate model, in terms of the higher forecasting accuracy and lower Type I and type two errors, Model I appears better than Model II and Model IIa in that

it has a higher prediction level (81.37% of correct classification) and very low Type II error (3.73%). However, the Type I error for Model I is quite high (14.9%). Since for international lenders it is potentially more important to correctly classify actual reschedulings (than actual non-reschedulings), that is have a lower Type I error, Model II (and IIa) is better in this context, generating a very low level of Type I error (3.73%). However, Model II and IIa have lower number of total correct predictions (75.78%) compared to Model I, which also implies that the Type II error has increased (20.5%). Such a Type II error of Model II (IIa) implies that the model is too restrictive, classifying 1/5 of non-rescheduling events as rescheduling. Thus applicability of this model is limited only for very risk averse investors. Further, Model II can keep many potential investors away from investing, as it overstates the possibility of rescheduling. This can decrease the growth possibilities in these small open economies slowing the speed of transition and increasing debt burden while waiting for potential investors. Therefore, the analysis suggests that in order to control both for the model accuracy and the Type I error, investors should find both of these models useful.

6. CONCLUDING REMARKS

This paper provides new empirical evidence about the determinants and importance of international and country specific macroeconomic indicators of the sovereign debt rescheduling probabilities of the EE countries. During the last decade, the EE countries became very attractive European region creating an investment boom. Specifically, the region became the largest recipient of net non-FDI flows among all emerging market regions in 2005. Developing models that can provide more accurate estimates of debt rescheduling probabilities in this region can benefit the investors as well as the host countries.

This study estimates sovereign default/rescheduling probabilities using panel logit models on a sample of 15 EE countries over the period 1990-2005. Compared to previous studies performed in this area, this paper combines different econometric methodologies to develop models with high degree of accuracy.

The findings show that in predicting the debt rescheduling probabilities one should utilise both of the two models suggested. Particularly, the models imply that EE countries, which want to reduce their probabilities of rescheduling, therefore reducing cost of borrowing and/or improve access to credit, should generally:

- 1) Control and reduce their General Government Expenditures by restructuring the public sector, attract more Foreign Direct Investments and stimulate GDP growth.
- 2) Increase their Export revenues by implementing a strategic trade policy in order to balance Current Account
- 3) Have a good repayment performance by limiting the size of the external debt compared to their GDP and do not default on sovereign debt as a result of unwillingness rather than inability to repay.

Furthermore, even though in the long run EU membership may improve the countries debt repayment performance and sovereign rating, our findings suggest that this political aspect is not significant for the sample period that we analysed.

APPENDIX 1

Table A1: Complete sample of sovereign debt rescheduling determinants: Descriptive Statistics

Variable	All Countries			Countries that have rescheduled			Countries that have not rescheduled		
	Obs	Mean	Std.dev	Ob	Mean	Std.dev.	Ob	Mean	Std.dev.
Lagged Rescheduling	176	0.239	0.427	35	0.7157	0.458	126	0.119	0.325
ICRG Rating Assigned	141	0.507	0.291	27	0.513	0.234	101	0.562	0.276
Total Debt/GDP	176	0.474	0.251	35	0.581	0.324	126	0.436	0.221
Total Debt/Exports	169	1.036	0.563	32	1.488	0.755	125	0.919	0.456
Short-term Debt/Total Debt	176	0.204	0.177	35	0.149	0.167	126	0.217	0.178
Interest Service due/Exports	169	0.041	0.027	32	0.036	0.019	125	0.0434	0.029
PNG, total private nonguaranteed/Exports	176	0.196	0.205	35	0.117	0.146	126	0.195	0.195
PPG, official creditors/Exports	176	0.425	0.480	35	0.852	0.557	126	0.315	0.390
PPG, total public and publicly guaranteed/Exports	176	0.754	0.630	35	1.365	0.707	126	0.608	0.517
Debt Service due/Exports	169	0.141	0.103	32	0.077	0.041	125	0.154	0.109
Reserves/Imports	168	0.035	0.014	31	0.030	0.016	125	0.036	0.012
Exports/GDP	176	0.431	0.179	35	0.317	0.113	126	0.458	0.180
Imports/GDP	176	0.509	0.174	35	0.384	0.152	126	0.536	0.166
Current Account Balance/GDP	173	-0.054	0.064	35	-0.028	0.091	125	-0.057	0.048
International Reserves/GDP	170	0.156	0.073	31	0.099	0.051	125	0.163	0.069
Credit to private sector/GDP	176	0.099	0.116	35	0.042	0.049	126	0.101	0.110
GDP per capita growth	175	0.033	0.058	34	0.020	0.089	126	0.033	0.048
GDP growth rate	175	0.028	0.051	34	0.012	0.065	126	0.029	0.048
Exports growth rate	168	0.095	0.140	33	0.072	0.227	120	0.095	0.113
Inflation rate (consumer prices)	174	0.515	1.774	35	1.328	3.306	124	0.343	1.077
Inflation rate (GDP deflator)	174	0.432	1.394	34	0.999	2.524	126	0.320	0.943
Devaluation of Exchange rate	174	-0.116	0.376	35	-0.284	0.534	124	-0.093	0.325
Interest arrears on LDOD/Exports	176	0.039	0.173	35	0.174	0.356	126	0.005	0.034
Principal arrears on	176	0.073	0.267	35	0.302	0.529	126	0.016	0.070
Interest arrears on LDOD/Debt	176	0.018	0.068	35	0.081	0.136	126	0.003	0.012
Principal arrears on LDOD/Debt	176	0.042	0.116	35	0.170	0.210	126	0.011	0.035
Domestic Saving Rate	176	0.153	0.131	35	0.123	0.201	126	0.163	0.104
Government Expenditure/GDP	173	0.179	0.048	34	0.183	0.048	124	0.178	0.049
M2/Total Reserves	169	0.031	0.021	31	0.042	0.021	124	0.029	0.020
S&P Emerging Market Index	85	0.195	0.559	6	0.539	1.265	69	0.115	0.471
Foreign Direct Investments/GDP	176	0.038	0.030	35	0.018	0.024	126	0.043	0.030
Real interest rate	151	0.037	0.164	29	0.034	0.278	109	0.038	0.131
Risk Premium on Landing	83	0.073	0.176	15	0.207	0.392	64	0.043	0.030
Stock Traded Total Value/GDP	134	0.044	0.063	23	0.029	0.045	98	0.045	0.065
Unemployment	127	0.129	0.060	20	0.141	0.057	103	0.126	0.060
Reserve/Total Debt	170	0.459	0.384	31	0.263	0.215	125	0.505	0.416
General government balance/GDP	122	-0.021	0.037	19	-0.032	0.046	98	-0.019	0.034
Government Revenue/GDP	123	13.380	45.061	19	74.312	82.763	99	0.291	0.112
Multilateral Debt/Total Debt	176	0.146	0.146	35	0.128	0.153	126	0.149	0.137
Concessional Debt/Total Debt	176	0.105	0.174	35	0.126	0.191	126	0.099	0.161

APENDIX 2 - Principal Component Analyses (PCA)**Table A2: Rotated Component Analysis Factor Matrix**

Rotated Factor Loadings – Verimax Rotation

Variable	1	2	3	4	5	Uniqueness
Dimension	Solvency	Liquidity	Currency Crises	Public Debt	External Debt	
Lagged rescheduling	-0.25901	0.70869*	-0.04018	-0.1655	0.41683	0.22793
ICRG rating	-0.422*	0.02938	-0.28118	0.06931	-0.37426	0.59711
Total debt/GDP	0.63688*	0.43336	-0.26262	0.06086	0.00979	0.33381
Short-term Debt/Total Debt	0.13345	-0.25657	0.08085	0.79881*	-0.09905	0.26191
Exports/GDP	0.77124*	-0.14041	0.00282	0.40539	-0.08846	0.21329
Imports/GDP	0.74658*	-0.18642	0.13958	0.37571	-0.32712	0.14022
Current Account Balance/GDP	-0.28283	0.1515	-0.12902	-0.05042	0.80626*	0.22782
International Reserves/GDP	0.7009*	-0.27745	0.30843	-0.06134	-0.00486	0.33285
Credit to private sector/GDP	0.68583*	-0.12632	0.22248	0.17689	-0.06644	0.42848
GDP growth rate	0.11705	-0.23265	0.75694*	0.09883	0.17676	0.3182
Inflation rate (consumer prices)	0.06741	0.00681	-0.79758*	-0.10676	0.29335	0.26183
Devaluation of Exchange rate	0.21231	-0.07999	0.87577*	0.06967	-0.06636	0.17229
Interest arrears on LDOD/Debt	-0.08535	0.88884*	-0.17785	0.03769	0.12409	0.15423
Principal arrears on LDOD/Debt	-0.1022	0.84736*	-0.00989	-0.14006	-0.06261	0.2479
Government Expenditure/GDP	0.2096	0.12429	0.12928	0.79534*	0.02363	0.29078
M2/Total Reserves	-0.41073	0.49661*	-0.4685	0.05535	-0.1585	0.337
Foreign Direct Investments/GDP	0.66853*	-0.27788	0.20131	-0.06331	-0.29768	0.3427

* the highest factor loading for each variable

REFERENCES

- Abiad, A., Leigh, D. and Mody, A. (2007), "International Finance and Income Convergence: Europe is Different", IMF Working Paper WP/07/64, IMF.
- Aylward, L., and Thorne R., (1998), "An Econometric Analysis of Countries' repayment Performance to the International Monetary Fund", IMF Working Paper No.32, March, International Monetary Fund
- Balkan, M. E., (1992), "Political instability, country risk and probability of default", *Applied Economics*, Vol. 24, pp. 999-1008
- Berg, A. and Sachs J.,D., (1988), "The Debt Crisis: Structural Explanations of Country Performance", *Journal of Development Economics*, Vol.29, pp. 271-306
- Brewer T.L., Rivoli P. (1997), Political instability and country risk, *Global Finance Journal*, Volume 8, Number 2, Autumn 1997, pp. 309-321(13)
- Butler, A., Fauver, L., (2006), "Institutional Environment and Sovereign Credit Ratings", *Financial Management*, Autumn, No. 35, 3, ABI/INFORM Global, pp. 53
- Camdessus, M., (1986), "Debt: Are we at the end of the crisis?", *European Economic Review*, No. 30, pp. 469-475.
- Carmen, A., (1992), "Debt Arrears in Latin America: Do Political Variables Matter?", *The Journal of Development Studies*, Vol. 28, No. 4, July, pp.668-688.
- Ciarlone, A., Trebeschi, G.,(2005) "Designing an early warning system for debt crises", *Emerging Markets Review*, No. 6., pp. 376-395
- Citron, J. and Nickelsburg, G., (1987), "Country Risk and Political Instability", *Journal of Development Economics*, Vol.25, pp.385-392

Coricelli, F., Mucci, F. and Revoltella, D., (2006), “Household Credit in the New Europe: Lending Boom or Sustainable Growth?” www.cepr.org/pubs/dps/DP5520.asp

Cuddington, J. and G. Smith, (1985), “International Debt and the Developing Countries”, Washington D.C.:World Bank.

Detragiache, E. and Spilimbergo, A., (2001) “Crises and Liquidity: Evidence and Interpretation”, IMF Working Paper, No. 2.

Eaton, J. and Gersovitz, J., (1981), “Debt with Potential Repudiation: Theoretical and Empirical Analysis”, *The Review of Economic Studies*, Vol. 48, No. 2. April, pp. 289-309.

Egert, B., Backe, P., Zumer, T., (2006), “Credit Growth in Central and Eastern Europe, New (over)shooting starts?”, European Central Bank, Working Paper Series, No. 687/October

European Central Bank, (2006), *Financial Stability Review*

European Commission (2006), “What do the sources and uses of funds tell us about credit growth in Central and Eastern Europe?”, European Economy, Occasional Papers, http://europa.eu.int/comm/economy_finance/index_en.htm

Feder, G. and Just, R., (1977),” A study of debt servicing capacity applying Logit analysis”, *Journal of Development Economics*, No. 4, pp. 25-38.

Feder, G., Just, R. E. and Knud, Ross, (1981), “Projecting Debt-Servicing Capacity of Developing Countries”, *Journal of Financial and Quantitative Analysis*, Vol.16, pp.651-669

Feder, G., and Uy, L.V. (1985), “The Determinants of International Creditworthiness and Their Policy Implications”, *Journal of Policy Modelling*, Vol.7, pp.133-156

Formatted: No bullets or numbering

Formatted: English (U.S.)

Frank, C. and Cline, W., (1971), "Measurement of debt servicing capacity: An application of discriminant analysis", *Journal of International Economics*, No. 1, pp. 327-344.

Georgievska, A., Georgievska, L., Stojanovic, A., Todorovic, N., (2009), "Sovereign rescheduling probabilities in emerging markets: a comparison with credit rating agencies' ratings", *Journal of Applied Statistics*, Vol.35, No.9, Sep 2008, 1031-1051

Haque, N., Nelson, M., Mathieson, D., (1998), "The Relative Importance of Political and Economic Variables in Creditworthiness Ratings", IMF Working Paper, No. 46.

Hajivassiliou, V., (1989), "Do the Secondary Markets Believe in Life After Debt?", World Bank Policy Planning and Research Working Paper No. 252, Washington, The World Bank

Hajivassiliou, V., (1994), "A Simulation Estimation Analysis of the External Debt Crises of Developing Countries", *Journal of Applied Econometrics*, Vol. 9, pp. 109-113

Hajivassiliou, V., (1987), "The External Debt repayment Problems of LDC's: An Econometric Model Based on Panel Data", *Journal of Econometrics*, Vol. 36, 205-230

Hamilton, L., C., (2004), *Statistics with Stata*, Thomson Learning

Hemming R. and Chalk N., (2000), "Assesing Fiscal Sustainability in Theory and Practice", IMF Working Paper, No. 81.

Hemming, R. and Petrie M., (2000), "A Framework for Assessing Fiscal Vulnerability", IMF Working Paper, No. 52

Hoti, S., (2005), "Modelling country spillover effects in country risk ratings", *Emerging Markets Review*, No. 6, pp. 324-345

Kiss, M., G. and Nagy, V., (2006), "Credit Growth in Central and Eastern Europe: Convergence or Boom?", Magyar Nemzeti Bank, MNB Working Papers 10

Kraay, A. and Nehru, V., (2004), “When Is External Debt Sustainable?”, World Bank Policy Research Working Paper 3200, February 2004.

Krugman Paul., (1987), “Financing vs. Forgiving a Debt Overhang”, *Journal of Development Economics*, No. 29., pp. 253-268.

Manasse, P., Roubini, N., Schimmelpfennig, A., (2003), “Predicting Sovereign Debt Crises”, IMF Working Paper, 221

McFadden, D., Eckaus, R., Feder, G., Hajivassiliou, V. and O’Connell S., (1985), “Is There Life After Debt? An Econometric Analysis of the Creditworthiness of Developing Countries”, in *International Debt and the Developing Countries*, edited by John Cuddington and Gordon Smith. Washington. The World Bank

Odedokun, M.O., (1995), “Analysis of Probability of External Debt Rescheduling in Sub-Saharan Africa”, *Scottish Journal of Political Economy*, Vol.42, pp.82-98

Reinhart, C. and Rogoff, K., (2004), “Serial Default and the "Paradox" of Rich-to-Poor Capital Flows”, *The American Economic Review*, Vol. 94, No. 2, pp. 53-58.

Roubini, N. and Menegatti, C., (2006), “Vulnerabilities in Central and Southern Europe”, www.rgemonitor.com

Sala-i-Martin, X (1997), “I just run four million regressions”, *The American Economic Review*, Vol. 87, No. 2, Papers and Proceedings of the Hundred and Fourth Annual Meeting of the American Economic Association, May, pp. 178-183

World Economic Outlook, (2006), “Emerging Europe: Managing Risks from Heavy Reliance on Foreign Savings”, Chapter II, IMF

World Bank EU8+2, (2007), Regular Economic Report Part II: Special Topic

World Economic Outlook, (2006), “Emerging Europe: Managing Risks from Heavy Reliance on Foreign Savings”, Chapter II, IMF

Table 1: Sample and Data

Country	Period examined	Rescheduling
Albania	1993-2005	5
Bulgaria	1992-2005	5
Bosnia & Hertzegovina	2000-2005	1
Czeck Republic	1994-2005	0
Estonia	1994-2005	0
Croatia	1994-2005	4
Hungary	1990-2005	0
Lithuania	1994-2005	0
Latvia	1994-2005	0
Macedonia	1997-2005	3
Poland	1991-2005	4
Romania	1991-2005	0
Russia	1995-2005	10
Serbia & Montenegro	2001-2005	3
Slovak Republic	1994-2005	0
Total	1990-2005	35

Table 2: Impact of selected variables on the probability of reheduling
Variables
Impact of the increase in value of the
variable on probability of reheduling
Positive **Negative**

Past rescheduling record:

Lagged Rescheduling

x

Political Variable:

ICRG Rating

x

Solvency variables:

Total Debt/GDP

x

Total Debt/Exports

x

Interest arrears on LDOD*/Exports

x

Principal arrears on LDOD*/Exports

x

Interest arrears on LDOD*/Debt

x

Principal arrears on LDOD*/Debt

x

Credit to private sector/GDP ^(a)

x

Real interest rate on international lending

x

Risk Premium on international landing

x

Exports/GDP

x

Exports growth rate

x

Multilateral Debt/Total Debt

x

Concessional Debt/Total Debt

x

Liquidity Variables:

Short-term Debt/Total Debt ^(b)

ambiguous

ambiguous

Interest Service due/Exports

x

Debt Service due/Exports

x

PNG**, total private nonguaranteed creditors/Exports

x

PPG***, official creditors/Exports

x

International Reserves/GDP

x

Reserve/Total Debt

x

Reserves/Imports

x

Variable used in currency crisis models:

Devaluation of Exchange rate

x

M2/Total Reserves

x

Macroeconomic variables:

Inflation rate (consumer prices)

x

Inflation rate (GDP deflator)

x

Unemployment

x

Foreign Direct Investments/GDP

x

GDP per capita growth

x

GDP growth rate

x

Domestic Saving Rate

x

Extrenal trade ratios:

Current Account Balance/GDP

x

Imports/GDP ^(c)

ambiguous

ambiguous

Public debt ratios:

Government Expenditure/GDP

x

General government balance/GDP

x

Government Revenue/GDP		x
PPG**, total public and publicly guaranteed debt/Exports	x	

Financial variables:

S&P Emerging Market Index		x
Stock Traded Total Value/GDP		x

Notes:

* Long-term debt outstanding

** Private non-guaranteed debt

*** Public and publicly guaranteed debt

- (a) If the private sector debt increases relative to the overall economy, the probability of rescheduling by private banks and companies increases. However, in transition countries, this variable can be the indicator of restructuring, development and future potentials for GDP growth, which has a negative relationship with the probability of sovereign debt rescheduling. Thus, indirectly, we expect this variable to be negatively related to probability of rescheduling.
- (b) The impact of this variable is ambiguous since on one hand the higher proportion of short term debt in total debt can lead to liquidity risks but on the other, investors can decrease their risk investing in country on short terms without being exposed to rescheduling probability in the long term period in future. Detragiache and Spilimbergo (2001) stress that in fact countries with imminent debt crises can only borrow short-term.
- (c) The higher the imports in relation to the size of the economy this country is more vulnerable to foreign shocks, and more likely to external debt rescheduling. Thus the expected coefficient of this variable should be positive (Frenkel, 1983). However, it can be debated whether this is always true. The higher this ratio, the more open economy is. Thus the country would not be willing to risk trade embargo or being ostracised in international economic arena due to default on external debt (Odedokun, 1995).

Table 3: Estimation Results

Dependent variable: rescheduling	<i>Model I</i>		<i>Model II</i>		<i>Model IIa</i>	
Independent Variables:	Coefficient	Marginal Effects %	Coefficient	Marginal Effects %	Coefficient	Marginal Effects %
ICRG rating			-.0176371	-.000071	-.0756585	-.0000398
Lagged rescheduling	-1.576432	-.0038672	3.013223 ***	.0385583	2.986011 ***	.005028
Total debt/GDP	7.694058 ***	.0255903	9.326913 ***	.0379242	9.290832 ***	.0048852
Exports/GDP	-21.16365 ***	-.0703899	-25.3023 ***	-.102882	-25.15219 ***	-.0132253
Current Account Balance/GDP	15.22603 **	.0506415	22.84288 **	.0928816	22.75349 **	.0119641
GDP growth			-13.61478 *	-.0553592	-13.46319 *	-.0070791
Foreign Direct Investments/GDP	-28.29895 *	-.0941218				
Government Expenditure/ GDP	57.87097 ***	.1924778				
Credit to private sector/GDP	-19.44009 **	-.0646574				
Short term debt/Total Debt	-6.047547 **	-.020114				
M2/Reserves			-24.15499	-.0982168	-23.79202	-.0125102
EU Dummy					-35.83584	-.0042394
constant	-5.396614**		2.16618		2.176749	
Sigma_u	4.949668		.0009119		.0009119	
Rho	.8816129		2.53e-07		2.53e-07	
Log-Likelihood	-36.385		-20.783		-20.753	
Wald Chi-Squared	17.40		19.32		18.92	
P-value of Chi2	0.0262		0.0072		0.0153	
Cut off point	0.5		0.5		0.5	
Correct Classifications(%)	81.37		75.78		75.78	
Type I Error (%)	14.91		3.73		3.73	
Type II Error (%)	3.73		20.5		20.5	
No. of Observations	157		120		120	
No. of Countries Analysed	15		12		12	
Period Analysed	1990-2005		1990-2005		1990-2005	

(a) Marginal effects are calculated at the sample means of the independent variables except for the dummy lagged rescheduling variable where the marginal effect is calculated for discrete change from 0 to 1

* significant at 10% level of significance

** significant at 5% level of significance

*** significant at 1% level of significance