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Benchmark-adjusted performance of US equity mutual funds and the issue of prospectus benchmarks

Abstract

This study examines the impact of mismatch between prospectus benchmark and fund objectives on benchmark-adjusted fund performance and ranking in a sample of 1281 US equity mutual funds. All funds in our sample report S&P500 index as a prospectus benchmark, yet 2/3 of those are placed in the Morningstar category with risk and objectives different to those of the S&P500 index. We identify more appropriate ‘category benchmarks’ for those mismatched funds, and obtain their benchmark-adjusted alphas using recent Angelidis et al. (2013) methodology. We find that S&P-adjusted alphas are higher than ‘category benchmark’-adjusted alphas in 61.2% of the cases. In terms of fund quartile rankings, 30% of winner funds lose that status when the prospectus benchmark is substituted with the one better matching their objectives. In the remaining performance quartiles there is no clear advantage of using S&P 500 as a benchmark. Hence, the prospectus benchmark can mislead investors about fund’s relative performance and ranking, so any reference to performance in a fund’s prospectus should be treated with caution.

Keywords: Prospectus benchmark selection, Mutual fund benchmark mismatch, Benchmark-adjusted alphas, Performance ranking

JEL: G11, G12, G23

1. Introduction

SEC regulations require mutual fund companies to disclose their performance relative to a passive benchmark, an index often referred to as their prospectus benchmark. Over a third of US investors rely on information in the fund prospectus when purchasing a mutual fund¹. Prospectus benchmark defines an investment direction and a risk tolerance, and should reflect the strategic role of the individual asset classes in the fund. However, Cremers and Petajisto (2009) provide evidence that mutual funds typically have a high proportion of holdings that differ from those of fund's (theoretically adequate) benchmark index. Sensoy (2009) affirms that funds frequently differ significantly from their benchmarks and shows that value funds are more likely to have self-designated benchmarks that are mismatched on value/growth, while small-cap funds tend to have prospectus benchmarks mismatched on size.

It should not come as a surprise then that some prospectus benchmark choices may be misleading, as there are no precise requirements on the selection of funds' best suited benchmark. Therefore, the choice of fund benchmark may be biased and may indicate principal-agent problems. As a consequence, for instance, a fund reporting a large cap index as their prospectus benchmark may have significant proportion of their assets invested in smaller size stocks. Considering investors' close scrutiny of fund performance it is vital to examine the extent of benchmark misclassification in US active fund management. Moreover, considering the development of recent literature on mutual fund performance, it is crucial to account for non-zero benchmark alphas, which significantly bias outcomes of fund performance (see for instance Chinthapati et al., 2017). A recent study by Cremers, Petajisto and Zitzewitz (2012) shows that standard benchmark models produce economically and statistically significant non-zero alphas for passive benchmark indices, including a widely used US passive benchmark - the S&P 500. Negative and statistically significant alpha for the Russell 2000 Growth index was documented by Chan, Dimmock, and Lakonishok (2009); significant non-zero alphas are also discussed in Costa and Jakob (2006). The non-zero alphas of passive benchmarks are not solely a US phenomenon.

¹ Investment Company Institute, Understanding Investor Preferences for Mutual Fund Information, Summary of Research Findings ("Understanding Investor Preferences"), 2006, available at https://www.ici.org/pdf/rpt_06_inv_prefs_full.pdf

Recently, persistent negative alphas are documented in the FTSE 100 Index in the UK (Mateus et al., 2016).

Based on the above, this paper aims to examine to which extent the benchmark choice of US long only equity funds changes inferences on fund performance, once the benchmark alphas are accounted for in fund performance evaluation. In particular, we assess whether inadequate prospectus benchmark selection may lead to over estimation of fund performance and whether it could be a subject of gaming. Further, we investigate whether benchmark choice affects fund performance in relative terms (relative to peers) and, therefore, changes the ranking position of the winning and losing funds, in particular. Hence, as our main contribution, we add to the literature on US mutual fund benchmark mismatch by 1) investigating the impact of the choice of benchmark on fund performance and performance rankings and 2) providing performance assessment free of biases caused by alphas embedded in the benchmark index and not accounted for in the standard pricing models. Recent literature suggests two methods to account for these non-zero benchmark alphas: Angelidis, Giamouridis and Tessaromatis (2013) and Chintlapati, Mateus and Todorovic (2017). In this paper, we apply Angelidis et al. (2013) methodology that adjusts the left hand side of the standard Carhart four-factor model, by replacing fund's risk adjusted return with the benchmark-adjusted return. As a result, this approach adjusts alpha of a fund by that of the fund's benchmark. Chintlapati, et al. (2017) follow the same intuition as Angelidis et al. (2013), whereby they use optimization algorithm² that calculates fixed minor adjustments for Carhart factors, which eliminate the alpha of a given benchmark index. Such adjusted factors are then used to estimate a fund's benchmark-adjusted alphas. They state that their method gives qualitatively the same results as Angelidis et al. (2013) approach, hence, in this paper we opt to use the latter method. Clearly, the choice of benchmark is critical for determining benchmark-adjusted alphas.

In the aspect of previous literature relevant to analysis, Sensoy (2009) provides evidence that funds frequently differ from their benchmarks in terms of their risk characteristics and composition for strategic reasons. Substantial exposures to size and value/growth factors in returns that are not captured by their benchmarks were also discussed in Elton, Gruber, and Blake (2003). The study of DiBartolomeo and Witkowski (1997) examine monthly returns

² The paper with Matlab code is available from: https://papers.ssrn.com/sol3/papers.cfm?abstract_id=2581737

for 748 load and no-load open-end funds and show that return patterns of 40 percent of funds analysed deviate from the benchmark declared in the prospectus with 9 percent of funds being seriously misclassified, two or more risk tiers away from their declared categories. Similarly, Kim, Shukla and Tomas (2000) assess how well mutual funds' stated objectives conform to their attributes-based objectives and revealed that the stated objectives of more than half the 1043 funds analysed differ from their attributes-based objectives, and over one third of the funds are severely misclassified. The study also confirms upward and downward risk shifts. Bams, Otten, and Ramezanifar (2016) analyse a sample of 1,866 US equity funds over the 2003-2015 period and found that 14% of funds are significantly misclassified based on long term style analysis. Huang et al. (2011) show that mutual funds change their total risk exposure substantially over time. Authors claim that it might be done for strategical reasons: in order to increase the expected money inflows to the funds or to manipulate their performance numbers. Similarly, Kacperczyk, Sialm, and Zheng (2008) measure the return gap, the difference between the reported fund return and the return on a portfolio that invests in the previously disclosed fund holdings, and document that despite disclosure requirements, mutual fund investors do not observe all actions of fund managers. Portfolio performance manipulation and deviation from benchmarks was also discussed in Goetzmann et al. (2007), Jiang et al. (2014), Fung and Hsieh (2002).

This paper contributes to the mutual fund performance measurement literature. In addition it adds to the literature on mutual fund benchmark misclassification and extends the work of Chan, Dimmock, and Lakonishok (2009), which demonstrates that judgments about the magnitude of performance are sensitive to the benchmarking methodology. We also extend the work of Sensoy (2009) by examining benchmark-adjusted performance and ranking using recently available Angelidis et al. (2013) methodology. To the best of our knowledge, this is the first study that analyses the impact of benchmark choice on US equity fund performance and ranking, while accounting for the non-zero alpha bias in the passive benchmarks. We use the net monthly returns of 1281 actively managed US equity mutual funds from January 1992 to February 2016. All funds in the sample report S&P500 as their primary prospectus benchmark in the Morningstar database. Our funds belong to 22 distinct Morningstar global categories: e.g. US Small Cap, US Large Cap Value, Energy Sector Equity, Global Equity etc. Investigation of commonly used benchmarks amongst funds in different categories in the Morningstar database, shows us that the primary prospectus benchmark that all our funds use, the S&P 500 Index, is most suitable for the funds in the Large Cap Blend Morningstar

category. However, around 2/3 of the funds in our sample are not in that category, nevertheless declaring S&P 500 as their primary benchmark. Our analysis of prospectus benchmarks fit shows that the funds' rationale for selecting a particular passive index as prospectus benchmark is not clear, as the index does not correspond to funds composition or investment objectives in large proportion of our sample. For each of the Morningstar global categories, we identify a benchmark index matching the objectives better than the S&P 500, which we refer to as the 'category benchmark' in this paper. We find that 'category benchmarks' are a better fit for our funds than their prospectus benchmark, the S&P 500 index, having on average around 10% higher R-squared in the full sample period and each of the sub-periods we examine. This makes an inference that even if fund alphas are adjusted for prospectus benchmark alphas, performance may be significantly biased if that benchmark is an unsuitable performance target.

To measure fund performance and provide rank of our funds, we apply Angelidis et al. (2013) methodology (AGT hereafter) that adjusts fund's alpha for benchmark's alpha, hence isolating manager's skill above that common to the benchmark. We find that 61.2% of the mutual fund AGT alphas are higher when S&P500 is used as a benchmark³. Further, in 15 out of 22 rolling periods of 36 months each, pairing the performance with S&P500 is beneficial to the funds and leads to overestimation of performance. Thus, on average, prospectus benchmark amplifies fund performance by 23 basis points versus the performance adjusted with a 'category benchmark'. This does not apply to all sub-periods, though. There is still 30% of sub-periods when benchmark-adjusted performance is better when the 'category benchmark' is used as the target in AGT model.

Analysis of fund quartile rankings shows that, on average, around 30% of winners leave the top quartile of funds when the benchmark is changed from the self-designated benchmark, S&P 500, to the 'category' benchmark in AGT benchmark-adjusted alpha estimation. On the opposite end of spectrum, nearly 30% of losers move up the quartiles when the 'category benchmark' is used. This finding supports the notion from Sensoy (2009) that funds appearing at the top end of the spectrum may choose their prospectus benchmarks strategically. In contrast, we find that inadequate prospectus benchmark actually harms the funds that are at the bottom of the ranks. Given this, we conclude that the choice of the

³ The results presented are obtained with the use of the Carhart model in AGT augmentation. The outcomes obtained with Fama-French three and five factor models are qualitatively the same and available upon request.

appropriate benchmark is critically important, as the wrong benchmark does not only bias performance assessment but can also lead to false conclusions when performance of funds relative to peers is assessed. Hence, this paper is of significance to retail investors, institutional investors and professional financial advisors interested in performance evaluation and fund rankings. Moreover, it has implications for financial regulators and policy makers with respect to fund information disclosure requirements and transparency in benchmark selection⁴.

This paper proceeds as follows: Section 2 describes the data. Section 3 provides preliminary analysis where we test the presence of alphas in passive benchmarks and evaluate which benchmark index is a best match to a fund's investment style. Section 4 presents the AGT methodology. Section 5 analyses funds' performance based on AGT-adjusted alphas and provides results. Section 6 looks at fund rankings and Section 7 concludes the paper.

2. Data

The data set comprises of 1,281 long-only active US equity mutual funds. The net monthly returns of mutual funds for the period January 1992 to February 2016 are from Morningstar (inclusive of dividends). A minimum of 36 months of returns is required for all the funds to be included in the sample. There is no survivorship bias. All funds in the sample declare S&P500 as their prospectus benchmark, but we observe that they follow variety of investment strategies, across all size and style categories as well as having sector or other focus.

In this paper, we use Morningstar global categories to facilitate the choice of an alternative index that would be a better-suited benchmark for fund's investment strategy than the self-declared S&P 500. The alternative benchmark can be determined using various methods/categorizations. For instance, one can estimate fund's sensitivity to Fama-French risk factors and derive the conclusions about the benchmark from factor loadings. However, this can be a biased approach as the factors are defined using arbitrary cut-offs for

⁴In the UK for instance, the Financial Conduct Authority (FCA) recently recognised the need for better transparency related to fund objectives and benchmark choice in their 'Asset Management Market Study'(published June 2017, accessed May 2018): <https://www.fca.org.uk/publication/market-studies/ms15-2-3.pdf>

constructing size and style portfolios (see Fama and French, 1993, Cremers, Petajisto and Zitzewitz, 2012, etc.) and can cause misclassification of (in particular large cap) funds (see Cremers et al. 2012 and Chen and Basset, 2014). Another alternative would be to use Sharpe (1992) style analysis, however, the model is found not capture well the sudden style drifts⁵. In contrast, Howard (2010) claims that funds should be grouped by their self-declared strategies rather than investment style box, and benchmark against such common strategy peer-group. However, peer group benchmarking is not the subject of this paper. Daniel, Grinblatt, Titman and Wermers (1997) suggest the use of benchmarks matching characteristics of stocks that mutual funds hold. Morningstar category allocation approach is in spirit similar to this, as Morningstar uses fund's historical holdings and portfolio statistics data to assign a fund to the category. Hence, in this paper, we use Morningstar global categories, which are commonly accepted as the industry-wide practice for fund classification and comparison⁶. I didn't provide a link to our lit review paper here, as I don't want to give referee ideas about what else to ask.

The list of Morningstar Global Categories where our funds are placed, the number of funds per each category, the most relevant passive benchmark for each category and the number of monthly observations per category are presented in Table 1. To select the most relevant passive benchmark for the category, we review the passive indices reported as benchmarks by all available funds in each category (not just those reporting S&P500 as their benchmark). In each category, we identify passive index most commonly reported as a benchmark, ensuring its characteristics correspond to the category it represents (e.g. US Large Cap Value category is best represented by Russell 1000 Value Index, which is also the most commonly reported benchmark in that category etc.). The returns data for all benchmarks is inclusive of dividends.

----Table 1---

Only 36% of our sample (460 funds) fall in the Large Cap Blend Morningstar Global category where the S&P500 would be deemed as the most appropriate passive benchmark. It means that performance analysis where the fund performance is measured against a prospectus benchmark can be biased and can provide inaccurate inferences about manager's skill. Further, 32% of our funds belong to the Large Cap Value and Large Cap Growth

⁵ <https://web.stanford.edu/~wfs Sharpe/art/fa/fa.htm> Accessed 28th September 2018.

⁶ There is no relevant change in categories of our funds over the sample period

Global category where most commonly reported benchmarks are Russell 1000 Value and Russell 1000 Growth index respectively. 112 out of 1,281 funds, or 8.7%, is in Midcap Global Category. Its better matched index would be Russell Midcap index, rather than S&P500. Further, some of our funds are in the Small Cap Morningstar Global category (40 of 1,281), best represented by a Russell 2000 index. Overall, these aforementioned five categories account 80 percent of our sample. All other Morningstar Global categories in our sample are sector specific or country/region specific and call for sector or regional benchmarks. These specialist funds account for the remaining 20% of our sample, Hence, significant proportion (64%) of our funds selects and reports a benchmark inappropriate for their category of funds. This is important from two perspectives: 1) measuring fund performance relative to the benchmark and 2) measuring fund performance relative to other similar funds. To this end, it is important to investigate fund's relative rankings within the same category and assess whether the funds that are the top performers according to prospectus benchmark (S&P500) change their relative ranking position after their performance is calculated with a benchmark that better reflects the risk characteristics of their Morningstar Global Category. Section 3.1 provides further discussion on suitability of the funds' self-declared benchmarks.

We split our analysis in 22 rolling overlapping sub-periods, each being 36 months of length. Given that the minimum data requirement for each fund is 36 months, within each rolling period we require that a fund has no less than 30 months of continuous returns. Table 2 reports the number of funds and monthly observations for each of the rolling sub-periods:

---Table 2---

3. Preliminary analysis

3.1 Test of the appropriateness of benchmark allocation

To begin with, we examine whether the 'category benchmarks' we have selected (as described in Section 2) provide a better fit than the self-declared prospectus benchmark, S&P500. To estimate this, we use the R-squared from equations (1) and (2) as a proxy for the accuracy of the benchmark used:

$$R_{i,t} = \alpha_{i,t} + \beta_i R_{S\&P500,t} + e_{it} \quad (1)$$

$$R_{i,t} = \alpha_{i,t} + \beta_i R_{\text{'category benchmark',t}} + e_{it} \quad (2)$$

In this analysis, we exclude the mutual funds that belong to the Large Cap Blend category (460 funds, as per Table 1) for which the ‘category benchmark’ (S&P500) is the same as their prospectus benchmark. For the remaining 821 funds we estimate equation (1) and (2), over the 22 rolling windows. Figure 1 depicts the average R-squared across the funds in each sub-period obtained using 1) the S&P as the benchmark and 2) ‘category’ benchmarks, as per equations (1) and (2).

---Figure 1---

The results confirm that the selected ‘category benchmarks’ are better suited than the S&P500 index for funds outside the Large Cap Blend category. The R-squared obtained using ‘category benchmarks’ for each sub-period and for the entire sample period is on average 10% higher, with peaks in 1999 and 2012, when the difference reached 14% and 11.5% respectively. The differences in R-squared are significant at 1% level in each rolling period and the overall sample (Wilcoxon z-test = 2.804)⁷. Given these results, the question that imposes itself is that of the impact of poorly suited benchmarks on the mutual fund performance and their ranking relative to other funds. Do funds with a prospectus benchmark unsuitable for their investment style tend to systematically outperform those benchmarks and do they remain at the top of the fund rankings when the benchmark is swapped for the more appropriate ‘category benchmark’? Before answering these questions, let us look into out/underperformance of the benchmarks themselves.

3.2 Presence of alphas in passive benchmarks

The issue we wish to avoid in our assessment of performance and ranks is that of the ‘closet-indexing’. For instance, if a ‘category benchmark’ (say, Russell 1000 Value) performs better than the self-designated benchmark chosen by a fund (S&P 500 here), the fund that belongs to that specific category (Large Cap Value in this example) is likely to outperform its self-reported benchmark (S&P 500), even if they are simply replicating their ‘category

⁷ The full set of R-squared values corresponding to Figure 1 and Wilcoxon z-tests of their difference are available on request.

benchmark' (Russell 1000 Value Index). If performance is assessed through standard Fama-French-Carhart framework, such funds may rank higher relative to other funds in the same category even though the fund managers exhibit no true skill.

To illustrate such bias inflicted by indices, in spirit of Costa and Jakob (2006), Chan, Dimmock, and Lakonishok (2009), Cremers, Petajisto and Zitzewitz (2012), Chinthalapari et al. (2017)⁸, we estimate the standard Carhart four-factor alphas of both self-declared prospectus benchmark (S&P500) and the 'category benchmarks' in our sample:

$$R_{Benchmark,t} - R_f = \alpha_{Benchmark} + \beta_M(R_{M,t} - R_{f,t}) + \beta_{SMB}SMB_t + \beta_{HML}HML_t + \beta_{WML}WML_t + e_t \quad (3)$$

Where $R_{Benchmark,t}$ is the return on the (prospectus or 'category') benchmark; R_f is the US 1 month Treasury bill; $R_{M,t} - R_{f,t}$ is the market risk premium⁹; SMB and HML are size and value factors from Fama and French (1993) paper and WML is the Carhart (1997) momentum factor. $\alpha_{Benchmark}$ represents the four-factor (prospectus or 'category') benchmark alpha, i.e. the excess return of the benchmark unexplained by the four factors.

The four-factor Carhart alpha is calculated for the S&P500, Russell 1000 Growth, Russell 1000 Value, Russell Midcap and Russell 2000 over 36 monthly rolling periods, to obtain alphas from 1994 to 2016. The aforementioned benchmarks correspond to the five largest Morningstar Global Categories in the data set and represent 80 percent of our fund sample (1,029 funds of a total 1,281). The remaining indices and their corresponding categories in our sample are not used for this analysis as the number of funds per category is not large enough resulting in some sub-periods featuring very few funds, jeopardising the objectivity of the results.

---Figure 2---

Figure 2 depicts the trend of annualized four-factor alphas (in bps) of the five indices. First, in line with previous studies (see for instance Chinthalapaty et al., 2017) the alphas of the five passive benchmarks are not zero. Specifically, the S&P500 and Russell 1000 Growth

⁸ who report non-zero alphas for passive benchmark indices

⁹ US market risk premium is defined as the value-weighted return of all CRSP firms incorporated in the US and listed on the NYSE, AMEX, or NASDAQ (R_m) minus one month US Treasury bill (R_f)

alphas tend to be more positive than those of the remaining indices analysed here. In the full sample period from January 1992 to February 2016, the S&P 500, Russell 1000 Growth and Russell Midcap indices all have positive four-factor alphas of 33.01, 74.93 and 60.17 basis points per year respectively; while the negative alphas of -12.58 and -197.01 basis points per year are obtained for the Russell 1000 Value and Russell 2000 index.

To get an indication of the magnitude of possible biases in fund performance evaluation by selecting an index not corresponding to funds' risk profile and composition holdings, we calculate the difference between the Carhart alpha of the 'category benchmark' and the self-declared benchmark, S&P500, as per Figure 3. The difference is annualized and reported in basis points.

---Figure 3---

Figure 3 illustrates that S&P500 four-factor alphas differ from the remaining four indices corresponding to the Global Categories where most of our funds 'reside'. For instance, in the sub-periods 1994-1996 and 1996-1998, the alpha for the self-declared prospectus benchmark S&P500 is positive but at least 100bps lower than the alpha for Russell 1000 Growth index. This tendency of the 'category' benchmark alpha to be higher than the prospectus benchmark one is present in 20 out of 22 rolling windows in this study. Therefore, it will be easier for a mutual fund in the Large Cap Growth category to outperform the prospectus benchmark (S&P500) as it has lower alpha relative to the 'category benchmark' one. If a fund "beats" the prospectus benchmark, investors may view that as a vouch for managerial skill, but it is possible that the fund is simply replicating Russell 1000 Growth, thus not having any stock picking skill. In that case, its outperformance over S&P500 should simply be attributed to a higher alpha of the 'category benchmark'. Note that this is not the case for all the indices. Inverse scenario can be noticed for Russell 2000, whose four-factor alpha is systematically lower than the S&P500 one. The differences between S&P500 and category indices' Carhart alphas are all significant at least at 5% level, with the exception of Russell Mid Cap Index. Wilcoxon z-stat for the difference between S&P500 and Russell 1000 Value is -2.098, Russell 1000 Growth is 3.142, Russell 2000 is -4.756 and Russell Mid Cap is 0.231.

To avoid the impact of 'closet indexers', there is a need to look at the benchmark-adjusted performance of funds. This is particularly important to note when measuring funds

performance and ranking relative to other funds. In the following section, we present the methodology that adjusts fund performance for benchmark performance and provides funds' benchmark-adjusted alphas.

4. Performance and ranking methodology

To obtain unbiased alphas for funds, we apply Angelidis, Giamouridis and Tessaromatis (2013) adjustment, suggested in recent literature on performance measurement¹⁰. The model is of interest to academics and investment professionals, as it adjusts the left hand side of the standard Carhart (1997) model by replacing the risk-adjusted return with the benchmark-adjusted return:

$$R_{i,t} - R_{Benchmark,t} = \alpha_i^* + \beta_{i1}^*(R_{M,t} - R_{f,t}) + \beta_{i2}^*SMB_t + \beta_{i3}^*HML_t + \beta_{i4}^*WML_t + e_i^* \quad (4)$$

where $R_{i,t} - R_{Benchmark,t}$ is the excess return of a mutual fund i over a benchmark in period t . As in equation (3) SMB and HML are size and value factors from Fama and French (1993) paper and WML is the Carhart (1997) momentum factor. All coefficients in this equation represent the difference between the coefficients of Carhart model performed on a fund and those of the Carhart model estimated on the benchmark index (equation 3). Thus, α_i^* is the difference of the fund's and benchmark's Carhart alpha (AGT-adjusted alpha hereafter). Similarly, coefficients $\beta_{i1}^*, \beta_{i2}^*, \beta_{i3}^*, \beta_{i4}^*$ represent the difference between a fund's and benchmark's Carhart betas (Angelidis et al, 2013, Mateus et al. 2016). If the coefficients $\beta_{i1}^*, \beta_{i2}^*, \beta_{i3}^*, \beta_{i4}^*$ are different from zero, this means that the mutual fund manager has different exposure to risk factors than the benchmark index. For instance, if the estimated SML AGT-adjusted beta (β_{i2}^*) is 0.15 it implies that the fund has 15% more exposure to small stocks than the benchmark. All of the factor data is from Kenneth French's website¹¹.

The AGT model, therefore, enables us to obtain AGT-adjusted four-factor alpha of a fund that accounts for the alpha of the benchmark. To assess the change in rankings when the benchmark changes from the prospectus benchmark (S&P500) to the 'category benchmark', the model will be used twice for each equity fund: with the S&P 500 as a benchmark and

¹⁰ Similar could be obtained using Chinthapati et al. (2017) methodology for benchmark-adjusted alphas

¹¹ Kenneth French's website: http://mba.tuck.dartmouth.edu/pages/faculty/ken.french/data_library.html

with the ‘category’ benchmark relevant for the Morningstar global category a fund belongs to:

$$R_{i,t} - R_{S\&P500,t} = \alpha_i^{*S\&P500} + \beta_{i1}^{*S\&P500}(R_{M,t} - R_{F,t}) + \beta_{i2}^{*S\&P500}SMB_t + \beta_{i3}^{*S\&P500}HML_t + \beta_{i4}^{*S\&P500}WML_t + e_i^{*S\&P500} \quad (5)$$

$$R_{i,t} - R_{Cat,t} = \alpha_i^{*Cat} + \beta_{i1}^{*Cat}(R_{M,t} - R_{F,t}) + \beta_{i2}^{*Cat}SMB_t + \beta_{i3}^{*Cat}HML_t + \beta_{i4}^{*Cat}WML_t + e_i^* \quad (6)$$

Where $R_{S\&P500,t}$ and $R_{Cat,t}$ are the return of the S&P 500 and ‘category’ benchmark respectively, $\alpha_i^{*S\&P500}$ is the difference between the Carhart alpha of fund i and its prospectus benchmark, S&P 500, α_i^{*Cat} is the difference between the Carhart alpha of fund i and the ‘category’ benchmark alpha; $\beta_{i1}^{*S\&P500}, \beta_{i2}^{*S\&P500}, \beta_{i3}^{*S\&P500}, \beta_{i4}^{*S\&P500}$ are fund i exposures to market risk, size, style and momentum factors beyond the exposure of S&P500 to those risks and $\beta_{i1}^{*Cat}, \beta_{i2}^{*Cat}, \beta_{i3}^{*Cat}, \beta_{i4}^{*Cat}$ are the fund i ’s four-factor betas adjusted by those of the ‘category’ benchmark for fund i ’s category. The rest is as per equation (4).

We estimate equations (5) and (6) for each fund and each of the 22 rolling sub-periods. Given that this is time-series analysis, we confirm that serial correlation and seasonality in residuals are not a cause for concern in our data sample¹². In total, we estimate 9,393 AGT S&P-adjusted alphas, $\alpha_i^{*S\&P500}$, and the same number of AGT ‘category benchmark’-adjusted alphas, α_i^{*Cat} . We rank the funds according to their $\alpha_i^{*S\&P500}$ in each of the 22 sub periods and split the funds in quartiles according to the performance. We do the same with α_i^{*Cat} to obtain the second set of fund quartile rankings.

5. Performance Results

Table 3 shows the number of unique funds analysed in each period, the average AGT S&P-adjusted alphas, $\alpha_i^{*S\&P500}$, and the average AGT ‘category benchmark’-adjusted alphas, α_i^{*Cat} , for each of the 22 overlapping periods from January 1992 to February 2016. All alphas are annualised averages across all categories, expressed in basis points. The table also reports the difference between $\alpha_i^{*S\&P500}$ and α_i^{*Cat} . In 15 over 22 periods the $\alpha_i^{*S\&P500}$ are higher

¹² Serial correlation of residuals does cause biases in our results. For 85% of the funds in our sample we accept the Breuch-Godfrey null hypothesis of no serial correlation with a lag of 1 (83% of funds when a lag of 12 is used to test for seasonality).

than the alternative (column 6 Table 3), implying that using S&P500 as a target instead of a more appropriate benchmark enhances performance. In some periods such as 2000-2002, 2003-2005 and 2006-2008 $\alpha_i^{*S\&P500}$ is higher than the α_i^{*Cat} in at least 79% of funds. In periods such as 2005-2007 and 2007-2009, the average $\alpha_i^{*S\&P500}$ exceeds average α_i^{*Cat} by over 200 basis points. However, this trend is not as pronounced in the period post-financial crisis: from 2009-2011 onwards, we find lower percentage of funds (e.g. 11% in 2011-2013) with an average $\alpha_i^{*S\&P500}$ higher than the alternative. In the full (non-overlapping) sample period, deploying the S&P 500 as the AGT adjustment instead of a ‘category benchmark’, overstates the performance for 61.2% of the funds. The difference in benchmark-adjusted AGT alphas stemming from the alternative benchmarks is significant in 50% of the rolling periods. For the overall sample, the difference in AGT adjusted alphas is significant at 1% level using Wilcoxon rank sum test ($Z = 5.75$). For robustness, we perform the same analysis using the AGT model based on the standard Fama-French three factor (Fama and French, 1993) and the relatively new, Fama-French five-factor model (Fama and French, 2015). The latter model is likely to become a new benchmark in asset pricing and fund performance measurement literature, in spite of its shortcomings¹³. We find that our conclusions associated with Table 3 still stand and that funds’ AGT alphas using category benchmarks are statistically significantly lower in the overall sample as well as in around 50% of rolling periods¹⁴.

---Table 3---

Although this evidence is pointing that using S&P500 as a benchmark in AGT model results in a better performance for a fund relative to the ‘category benchmark’ in most of the rolling sub-periods, we do not know whether this benefits more the funds at the top or at the bottom of performance ranks. One should not ignore the fact that there is still 38.8% of the funds in our sample that are worse off by indicating S&P500 as a prospectus benchmark. To further examine the issue of strategic benchmark choices, we investigate whether the fund rankings change considerably when the prospectus benchmark is replaced with a relevant ‘category’ one.

¹³ The factors in the five-factor model are market, size, style, investment and profitability. For shortcomings of the Fama-French five-factor model see for instance Fama and French (2016) and Blitz, Hanauer, Vidojevic and van Vliet (2018)

¹⁴ Full set of these results is available on request.

6. The impact of benchmark choice on fund rankings

We examine how the choice of benchmark may impact funds' relative ranking: do winners tend to stay winners and do losers remain losers when the benchmark changes from the one disclosed in the prospectus (S&P 500) to the 'category benchmark'. Using the AGT-adjusted alphas for each fund over 22 rolling periods, we split the funds into quartiles in each period. Two sets of quartile rankings are constructed one based on AGT S&P500-adjusted alphas and one on AGT 'category benchmark'-adjusted alphas. Quartile ranking is not done in each Morningstar Global Category but overall, as some categories have small number of funds in a number of sub-periods. We construct quartiles using the funds in all the categories excluding those assigned to the Large Cap Blend Global category, as their 'category benchmark' is their prospectus benchmark, the S&P500 index.

We then examine the proportion/number of funds that change quartiles when the benchmark changes. Table 4 displays the number of funds in each quartile per rolling period, the average annualised AGT alphas (in bps) adjusted for i) prospectus, S&P500 ($\alpha_i^{*S\&P500}$) and ii) 'category benchmark' (α_i^{*Cat}) per each rolling period. The table also reports the difference between the two alphas, which signals the magnitude of a possible bias when inappropriate benchmark is used in performance assessment. The last column in Table 4 shows the percentage of funds that remain in the same performance quartile when S&P 500 index is replaced with the relevant 'category benchmark'.

---Table 4---

In Panel A (Quartile 1) 'category benchmarks' provide a lower average AGT adjusted alpha in 12 out of 22 periods analysed, indicating that for 55% of the periods performance of winners estimated with S&P500 is overstated. Analogous tendency can be viewed for the Quartile 2, 3 and 4 (11, 13 and 11 out of 22 periods of lower average AGT 'category benchmark'-adjusted alphas, α_i^{*Cat} , respectively).

More importantly, the average number of funds that remains in the top quarter over the years (70%), implies that 30% of the top performing funds drop in performance ranks and leave the

quartile when the performance is adjusted with the ‘category benchmark’. Overall, for the total sample period, there is on average 68bps advantage for Quartile 1 funds of using S&P 500 as the prospectus benchmark. Comparing this value to the equivalent average alpha difference in Panels B-D, it becomes evident that the top performing funds benefit the most from the choice of prospectus benchmark. Panel D in fact suggests that Quartile 4 funds get penalised by inadequate benchmark selection. Thus, on average, close to 30% of loser funds move up in quartile rankings when their performance is assessed against a ‘category benchmark’. The average AGT ‘category benchmark’-adjusted alpha of loser funds for the total period is 33bp higher than the one estimated with prospectus benchmark, leading us to conclude that these funds would be better off selecting a relevant ‘category benchmark’.

Quartile 2 and Quartile 3 funds (Panel B and C of Table 4) are of least interest to investors; the funds in these quartiles are neither the top funds investors look out for nor the ones at the bottom they are trying to avoid. However, we document that the results for both quartiles are similar: adjusting alphas with the relevant ‘category benchmark’ changes, on average, the quartile ranking of 45% and 43% of funds from Quartile 2 and Quartile 3, respectively. Those movements can be in both directions – up to a higher or down to a lower ranked quartile, and in most of the cases there is an interchange between these two groups. The AGT alpha adjusted with the ‘category benchmark’ is on average 28 (Quartile 2) and 26 (Quartile 3) basis points higher than the one adjusted with self-declared prospectus benchmark.

For robustness and comparison, we replicate the analysis using standard Fama-French three-factor and Fama-French five-factor model versions of the AGT, with the S&P 500 index and the category benchmark. Table 5 shows that our results remain qualitatively the same and quantitatively very similar¹⁵. When we adjust for the category benchmark, around 30% of the funds drops out of the top quartile of performance regardless of the number of factors used in the AGT model. Similarly, around 30% of funds leaves the bottom quartile, regardless of the model used.

-Insert Table 5-

¹⁵ The full set of tables for AGT with three-and five-factors, fully comparable to Table 4 based on four-factor model are available on request.

Therefore, inferences on mutual fund relative performance may be significantly biased when fund performance is evaluated in respect to unsuitable benchmark. To support our discussion, we plot the difference in average AGT adjusted alphas, $\alpha_i^{S\&P500} - \alpha_i^{Cat}$ in each ranking quartile and each period (column 5 from Table 4) in Figure 4.

---Figure 4 ---

The figure shows that the average AGT-adjusted alphas for the Quartile 2 and 3 are almost identical irrespective of the benchmark. However, the performance of top funds is overestimated with the prospectus benchmark in over half of the rolling periods. The difference in AGT adjusted alphas for Quartile 1 funds reaches peaks of -241bps in 1993-1995, around -300 bps in 2006-2008 and 2008-2010, and a maximum of -460bps as in 2007-2009, in favour of alphas adjusted with the S&P rather than the ‘category benchmark’. Even though top performing funds seem to take advantage of using S&P500 as their benchmark; there are also cases when performance of these funds benchmarking against prospectus benchmark could be undervalued, as in 2000-2002 and 2001-2003, but by a smaller margin. In contrast, our results show, that while benchmarking against prospectus benchmark is on the whole beneficial to winners, it negatively affects the performance of losers. The difference in AGT adjusted alphas of Quartile 4 funds in some time periods, for instance 2011-2013 and 2012-2014, reaches 392 and 309 basis points, respectively, in favour of alphas adjusted with the ‘category benchmark’.

Considering these findings, most of the funds that are potentially strategically selecting S&P 500 as the benchmark and benefiting from it are those in top performance quartile. They have on average 0.68% higher benchmark-adjusted alphas when that benchmark is the one given in the prospectus and nearly 30% of those funds lose the ‘winner’ status when the self-declared benchmark is substituted with a better suited one. In all other quartiles there is no clear advantage of using S&P 500 as a prospectus benchmark. Hence, the choice of the benchmark affects not only the inferences about a fund’s absolute performance, but it can also mislead investors about its relative performance. This leads us to conclude that any information in fund prospectus about the performance relative to the prospectus benchmark or relative to other funds should be treated with caution.

7. Conclusions

This paper analyses the impact of benchmark choice on US equity funds performance and ranking and gauges potential biases in absolute and relative performance evaluation stemming from the inaccurate prospectus benchmark selection. We evaluate the question of mismatch between the prospectus benchmark and fund objectives, raised in Sensoy (2009), and estimate the impact of such misclassification on fund performance and ranking when recently available benchmark-adjusted performance measure is used. Hence, our analysis accounts for non-zero benchmark alphas produced by standard pricing models, discussed in recent literature such as Chan et al. (2009), Cremers et al. (2012) and Chinthlapati et al. (2017). Our sample includes net monthly returns of 1281 actively managed US equity mutual funds from January 1992 to February 2016 reporting S&P 500 index as their primary prospectus benchmark in the Morningstar database. We find that only 460 of those funds belong to the Large Cap blend Morningstar category, for which the S&P 500 would be the most suited benchmark. All other remaining funds fall across 21 other distinct Morningstar Global categories, some of which imply that fund risk profile and composition is very different from that of their prospectus benchmark. Naturally, we investigate whether the fund's performance relative to the S&P 500 is better than when measured against what we consider their relevant 'category' benchmark. Regression of mutual fund returns on the returns of S&P500 and the benchmark relevant for the Morningstar global category a fund belongs to, shows that the 'category' benchmarks are a better fit for our funds, having on average 10% higher R-squared.

Further, in our preliminary analysis, we report non-zero alphas of passive benchmark indices in our sample. To eliminate the upward/downward biases in performance assessment caused by embedded benchmark alphas, we apply Angelidis et al. (2013) method (AGT) that adjusts a fund's alpha for benchmark's alpha, hence isolating manager's skill above that common to the benchmark. Performance for each fund is calculated against the S&P500 index and the 'category benchmark', more appropriate for the Morningstar category a mutual fund belongs to. The sample period is split into 22 rolling overlapping windows, each being 36 month in length. In the total sample period we document higher AGT four-factors alphas estimated with S&P 500 as a benchmark versus those adjusted with the relevant 'category benchmark'. In 70 percent of the sub-periods the average AGT alphas adjusted with S&P500 are higher than those adjusted with 'category benchmark', thus overestimating fund performance.

Overall figures for the entire period show that 61.2% of the funds benefit from wrongly benchmarking their performance against the prospectus benchmark, S&P500; the average AGT-adjusted alpha drops by 23 basis points when relevant ‘category benchmark’ is used.

Additional results show that poor benchmark choice also influences relative performance assessment. We find that the top quartile funds benefit most from the choice of prospectus benchmark. For instance, in 2007-2009 the difference in S&P 500- and ‘category benchmark’-adjusted alphas reached 460 bps in favour of using the prospectus benchmark. Also, 30% of top performing funds move their ranking position downwards when their performance is adjusted with the ‘category benchmark’ instead of the S&P500. The results also show that the worst performing funds get penalised by their prospectus benchmark choice. In fact, close to 30% of losers move up the quartiles when performance is estimated with the most suitable ‘category benchmark’. This leads us to conclude that strategic benchmark selection appears to be most likely in the funds at the top performance quartile, while we do not observe clear advantage of benchmark gaming in the remaining quartiles.

Our paper shows that appropriate benchmarking is essential for accurate performance evaluation, as inferences on both fund performance and performance ranking may change significantly when estimated against a ‘category benchmark’ instead of their self-declared prospectus benchmark. This study raises concerns that require attention of financial regulators and policy makers. New information disclosure requirements should be placed to provide more clarity for investors as to how the prospectus benchmark is selected. It also calls for investors to be more cautious when interpreting performance figures in fund prospectus. The paper can be extended to non-equity funds or funds where the benchmarking is ambiguous (such as hedge funds for instance).

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Figure 1: Average R-squared of S&P500 and Global 'Category benchmark' fit

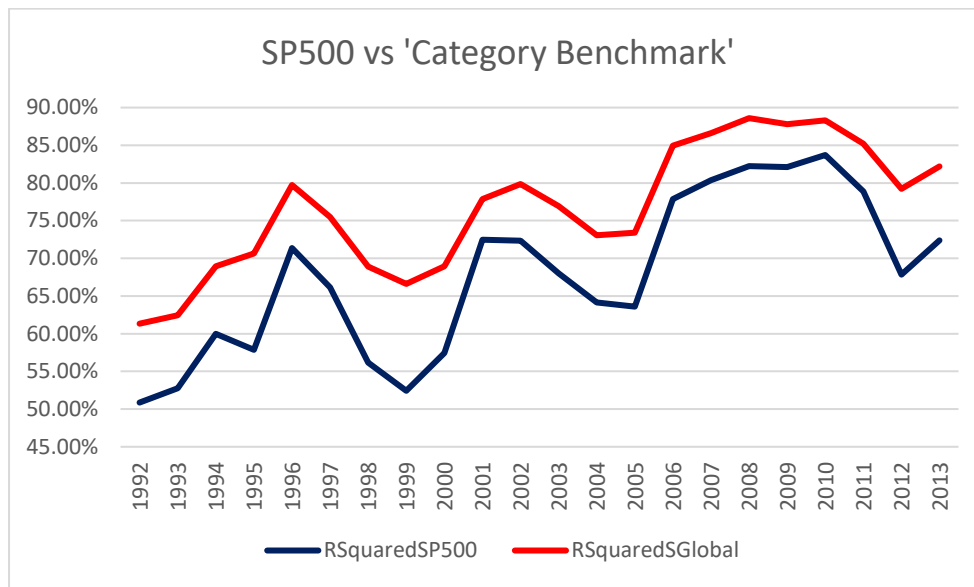


Figure 2: Four factor (Carhart) alphas of S&P 500 and selected ‘Category’ benchmarks

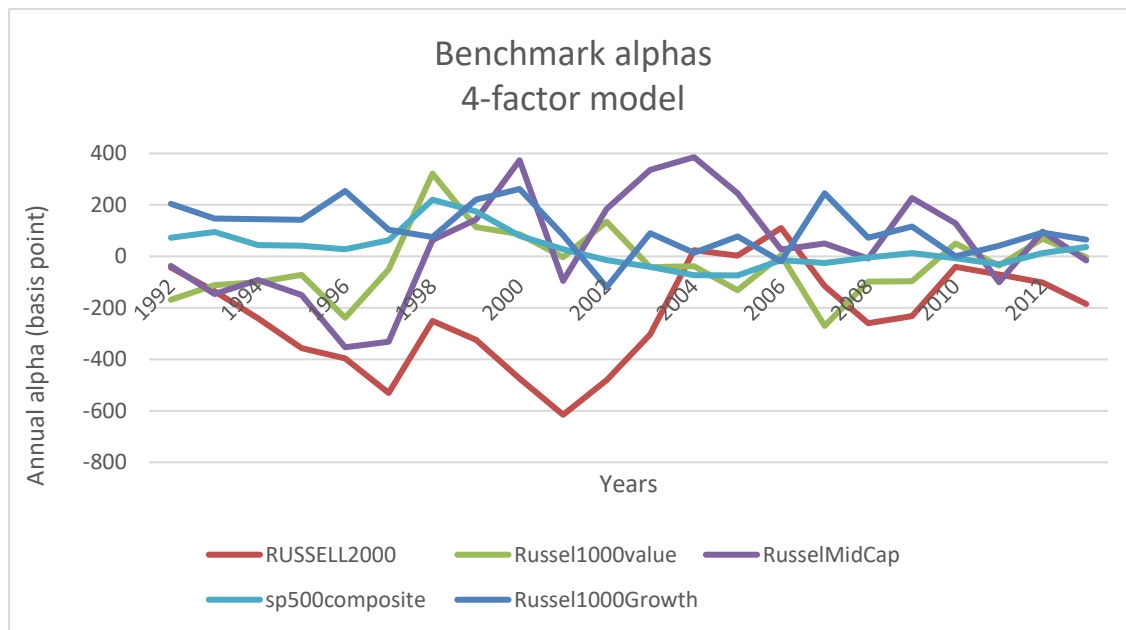


Figure 3 Differences between Carhart alphas of relevant ‘Category’ benchmarks and the S&P500

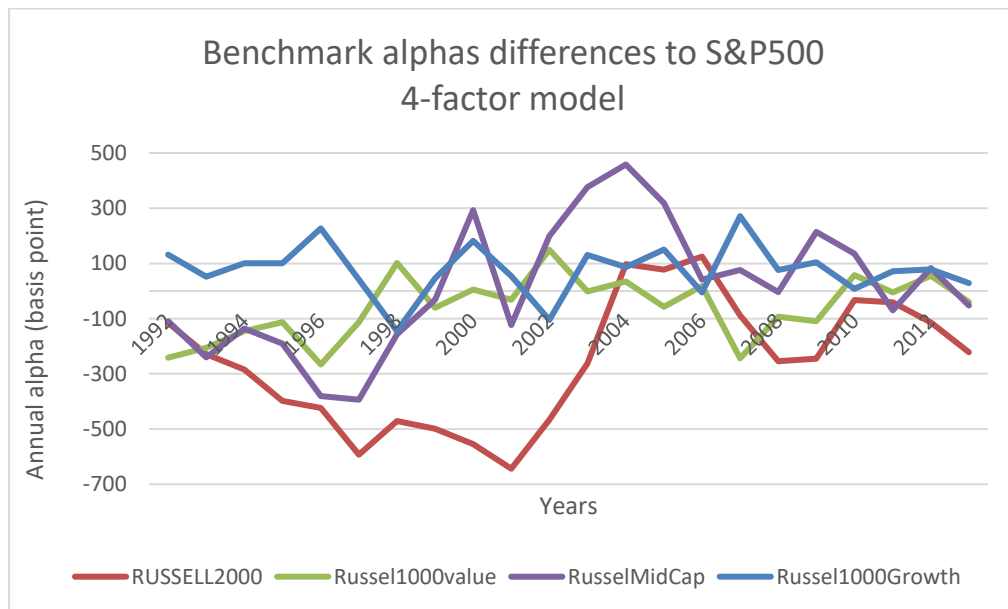


Figure 4: Difference between AGT S&P adjusted and AGT 'category' benchmark adjusted alphas

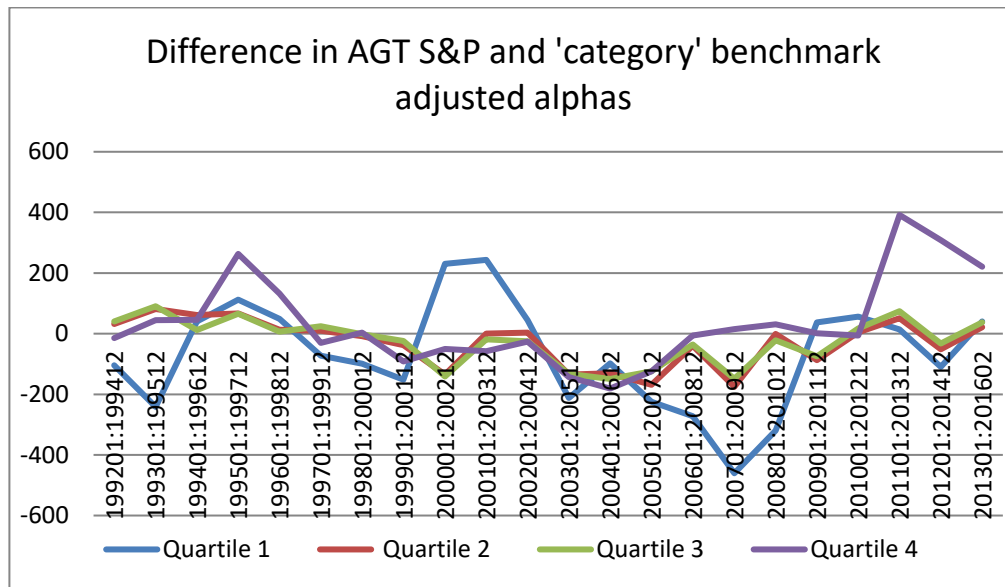


Table 1: Sample of ‘category’ benchmarks

The sample consists of 1,281 (212,122 monthly observations) long-only active US equity mutual funds from January 1992 to February 2016. For all funds the self-declared prospectus benchmark is the S&P500. Table below shows the Morningstar Global Category our funds belong to, the suitable benchmark for the category, the number of funds in the category and number of monthly observations per category (all benchmarks are total return and in USD).

Global Category	Suitable Benchmark	# Funds	# Monthly Observations
US Large Cap Blend	S&P 500	460	73,493
US Large Cap Growth	RUSSELL 1000 GROWTH	290	48,393
US Large Cap Value	RUSSELL 1000 VALUE	127	21,160
US Mid Cap	RUSSELL MIDCAP	112	17,332
Technology Sector Equity	S&P500 ES INFO TECHNOLOGY	54	9,092
US Small Cap	RUSSELL 2000	40	5,611
Healthcare Equity	S&P500 ES HEALTH CARE	32	5,554
Real Estate Equity	S&P500 DIVERSIFIED REIT'S	24	2,279
Global Equity	MSCI WORLD	22	3,392
Financial Sectors Equity	S&P500 DIVERSIFIED FINANCIALS	19	4,162
Energy Sector	S&P500 ENERGY IG	16	3,198
Precious Metals Sector Equity	S&P GSCI Precious Metal Tot. Ret.	16	4,196
Utilities Sector	S&P500 ES UTILITIES	14	3,293
Natural Resources Equity	S&P GSSI NORTH AMER. NAT.RES.SECTOR	13	2,400
Consumer Goods and Services	S&P500 ES CONSUMER DISCRETIONARY	12	3,159
Industrials Equity	S&P500 ES INDUSTRIALS	8	2,124
Communications Equity	S&P500 COMM. EQUIPMENT	8	1,268
Global Equity Large Cap	MSCI EAFE	7	1,174
Emerging Markets Equity	MSCI EM	2	324
Other Equity (Emerging Europe)	MSCI EM EUROPE	1	227
Europe Large Cap Equity	MSCI EUROPE	1	82
Asia Equity ex Japan	MSCI AC ASIA PAC EX JP	1	58
Japan Equity	MSCI JAPAN	1	61
Greater China	MSCI GOLDEN DRAGON	1	90
Total:		1,281	212,122

Table 2: Sample funds with more than 36 monthly observations

Table reports the number of funds and monthly observations for each of the 36 months rolling windows. The minimum data requirement is for funds to have at least 36 months of continuous observations. The #Funds represents the number of (non-unique) funds with available data in each period.

<i>Period</i>	<i># Funds</i>	<i># Monthly Observations</i>	<i>Period</i>	<i># Funds</i>	<i># Monthly Observations</i>
<i>199201:199412</i>	409	12,508	<i>200301:200512</i>	1,034	32,887
<i>199301:199512</i>	451	14,042	<i>200401:200612</i>	1,070	33,361
<i>199401:199612</i>	527	15,740	<i>200501:200712</i>	1,066	33,956
<i>199501:199712</i>	600	17,860	<i>200601:200812</i>	1,054	34,366
<i>199601:199812</i>	681	20,463	<i>200701:200912</i>	1,057	33,663
<i>199701:199912</i>	771	23,364	<i>200801:201012</i>	1,039	32,453
<i>199801:200012</i>	865	26,305	<i>200901:201112</i>	975	30,906
<i>199901:200112</i>	919	28,916	<i>201001:201212</i>	895	29,500
<i>200001:200212</i>	955	30,874	<i>201101:201312</i>	855	27,929
<i>200101:200312</i>	980	32,085	<i>201201:201412</i>	789	26,519
<i>200201:200412</i>	997	32,640	<i>201301:201602</i>	751	26,573
<i>Overall: 199201:201602</i>				1,281	211,855

Table 3: Comparison of average AGT S&P adjusted alphas and average AGT ‘category’ benchmark adjusted alphas

The table reports comparison of alphas from the following two regressions:

$R_{i,t} - R_{S\&P500,t} = \alpha_i^{*S\&P500} + \beta_{i1}^{*S\&P500}(R_{M,t} - R_{F,t}) + \beta_{i2}^{*S\&P500}SMB_t + \beta_{i3}^{*S\&P500}HML_t + \beta_{i4}^{*S\&P500}WML_t + e_i^{*S\&P500}$ and $R_{i,t} - R_{Cat,t} = \alpha_i^{*Cat} + \beta_{i1}^{*Cat}(R_{M,t} - R_{F,t}) + \beta_{i2}^{*Cat}SMB_t + \beta_{i3}^{*Cat}HML_t + \beta_{i4}^{*Cat}WML_t + e_i^{*Cat}$. $R_{S\&P500,t}$ and $R_{Cat,t}$ are the return of the S&P 500 and ‘category’ benchmark relevant for Morningstar Global Category respectively, $\alpha_i^{*S\&P500}$ is the difference between the Carhart alpha of fund i and its prospectus benchmark, S&P 500, α_i^{*Cat} is the difference between the Carhart alpha of fund i and the ‘category’ benchmark alpha; $\beta_{i1}^{*S\&P500}, \beta_{i2}^{*S\&P500}, \beta_{i3}^{*S\&P500}, \beta_{i4}^{*S\&P500}$ are fund i exposures to market risk, size, style and momentum factors beyond the exposure of S&P500 to those risks and $\beta_{i1}^{*Cat}, \beta_{i2}^{*Cat}, \beta_{i3}^{*Cat}, \beta_{i4}^{*Cat}$ are the fund i ’s four-factor betas adjusted by those of the ‘category’ benchmark relevant for fund i ’s category. $R_{M,t} - R_{F,t}$ is the market risk premium; SMB and HML are size and value factors from Fama and French (1993) paper and WML is the Carhart (1997) momentum factor. Alphas and the difference in alphas are annualized and given in bps. ***, **, * denote significance at 1%, 5% and 10% level based on Wilcoxon z-stat.

<i>Period</i>	<i># of funds</i>	<i>Average</i> $\alpha_i^{*S\&P500}$ <i>(bp)</i>	<i>Average</i> α_i^{*Cat} <i>(bp)</i>	<i>Average difference</i> $\alpha_i^{*Cat} - \alpha_i^{*S\&P500}$ <i>(bp)</i>	<i>Better</i> $\alpha_i^{*S\&P500}$ <i>#/%</i>	<i>Better</i> α_i^{*Cat} <i>#/%</i>
199201:199412	192	43	4	-39	110/58%	81/42%
199301:199512	218	28	-12	-40	121/56%	97/44%
199401:199612	245	-208	-151	57*	113/46%	132/54%
199501:199712	275	-356	-217	139*	124/45%	151/55%
199601:199812	299	-274	-261	14	143/48%	156/52%
199701:199912	344	-318	-363	-45	165/48%	179/52%
199801:200012	384	341	214	-127	261/68%	123/32%
199901:200112	433	279	198	-81	310/72%	123/28%
200001:200212	476	47	-14	-61***	392/82%	84/18%
200101:200312	526	-244	-217	27	311/59%	215/41%
200201:200412	534	-154	-156	-2	287/54%	247/46%
200301:200512	514	29	-111	-140***	426/83%	88/17%
200401:200612	524	8	-85	-93***	402/77%	122/23%
200501:200712	514	333	78	-255***	384/75%	130/25%
200601:200812	513	113	-17	-130***	405/79%	108/21%
200701:200912	513	228	2	-226***	382/75%	131/25%
200801:201012	506	87	-32	-119**	345/68%	161/32%
200901:201112	490	-9	120	129***	330/67%	160/33%
201001:201212	487	-225	-178	47	201/41%	286/59%
201101:201312	484	-361	-154	207***	55/11%	429/89%
201201:201412	473	-313	-183	130	200/42%	273/58%
201301:201602	449	-330	-231	99	280/62%	169/38%
Overall		-49	-90	-41***	5,747/61.2%	3,645/38.80%

Table 4: Difference in alphas per quartile and change of quartile ranks

Panels A-D report results for Quartile 1(top) - 4 (bottom) respectively. All panels show the number of funds and comparison of AGT adjusted alphas, when S&P 500 is used as a benchmark ($\alpha_i^{*S\&P500}$ from eq. (5)) and when 'category' benchmark is used (α_i^{*Cat} from equation (6)). Alphas and the difference in alphas are annualised and given in basis points. The last column shows percentage of funds that remains in the same quartile when the benchmark is changed from the S&P500 to the relevant 'category' benchmark. In the last row, the 'average' represents the average across the periods and across the funds.

Panel A: Quartile 1 (Carhart model)					
<i>Period</i>	<i># of Funds</i>	<i>Average</i> $\alpha_i^{*S\&P500}$ <i>(bp)</i>	<i>Average</i> α_i^{*Cat} <i>(bp)</i>	<i>Average difference</i> $\alpha_i^{*Cat} - \alpha_i^{*S\&P500}$ <i>(bp)</i>	<i>% Funds</i> <i>remaining in</i> <i>Quartile 1</i>
199201:199412	48	724.622	619.908	-104.713	77.08
199301:199512	55	900.993	659.124	-241.869	78.18
199401:199612	61	509.698	550.235	40.537	72.13
199501:199712	69	406.961	518.888	111.927	71.01
199601:199812	75	401.890	450.835	48.945	50.67
199701:199912	86	710.099	637.692	-72.407	75.58
199801:200012	96	1513.924	1414.484	-99.440	68.75
199901:200112	108	1431.359	1279.641	-151.718	76.85
200001:200212	119	1079.855	1310.632	230.777	64.71
200101:200312	132	744.157	987.541	243.384	68.94
200201:200412	134	477.240	521.377	44.137	65.67
200301:200512	129	649.942	438.095	-211.847	72.09
200401:200612	131	643.598	545.615	-97.982	78.63
200501:200712	129	667.274	442.562	-224.712	68.22
200601:200812	128	675.317	402.982	-272.335	75.00
200701:200912	128	1049.981	589.334	-460.646	58.59
200801:201012	127	876.886	556.633	-320.253	77.95
200901:201112	123	586.092	623.681	37.589	61.79
201001:201212	122	390.537	446.654	56.1171	68.03
201101:201312	121	406.372	420.2614	13.889	71.07
201201:201412	118	496.183	386.647	-109.535	69.49
201301:201602	112	234.353	274.290	39.937	71.43
Average				-68.19	70.09

Panel B: Quartile 2 (Carhart model)					
<i>Period</i>	<i># of Funds</i>	<i>Average</i> $\alpha_i^{S\&P500}$ (bp)	<i>Average</i> α_i^{Cat} (bp)	<i>Average difference</i> $\alpha_i^{Cat} - \alpha_i^{S\&P500}$ (bp)	<i>% Funds</i> <i>remaining in</i> <i>Quartile 2</i>
199201:199412	48	66.766	98.823	32.057	43.75
199301:199512	54	20.647	101.976	81.329	53.70
199401:199612	61	-64.974	-3.744	61.229	52.46
199501:199712	69	-47.359	19.710	67.069	44.93
199601:199812	75	-83.858	-70.943	12.916	9.33
199701:199912	86	-103.845	-93.813	10.033	61.63
199801:200012	96	285.842	277.472	-8.370	54.17
199901:200112	108	284.427	246.554	-37.873	64.81
200001:200212	119	261.234	126.598	-134.636	64.71
200101:200312	131	-93.440	-93.051	0.389	64.89
200201:200412	133	-22.030	-19.082	2.948	45.11
200301:200512	128	117.070	-17.686	-134.756	49.22
200401:200612	131	197.198	66.072	-131.125	54.96
200501:200712	128	242.302	73.913	-168.389	57.03
200601:200812	128	84.538	44.437	-40.100	71.09
200701:200912	128	261.924	86.677	-175.248	32.03
200801:201012	126	47.480	46.723	-0.757	65.87
200901:201112	122	22.702	-65.162	-87.864	63.11
201001:201212	122	-17.213	-13.158	4.055	61.48
201101:201312	121	-37.196	13.985	51.181	71.07
201201:201412	118	-20.476	-72.422	-51.946	67.80
201301:201602	112	-121.013	-99.772	21.241	67.86
Average				-28.48	55.50

Panel C: Quartile 3 (Carhart model)					
<i>Period</i>	<i># of Funds</i>	<i>Average</i> $\alpha_i^{*S\&P500}$ <i>(bp)</i>	<i>Average</i> α_i^{*Cat} <i>(bp)</i>	<i>Average difference</i> $\alpha_i^{*Cat} - \alpha_i^{*S\&P500}$ <i>(bp)</i>	<i>% Funds</i> <i>remaining in</i> <i>Quartile 3</i>
199201:199412	47	-187.715	-147.318	40.397	38.30
199301:199512	54	-224.299	-133.920	90.379	51.85
199401:199612	62	-306.194	-294.377	11.817	50.00
199501:199712	68	-352.285	-286.392	65.893	52.94
199601:199812	74	-357.022	-350.239	6.783	25.68
199701:199912	86	-402.512	-377.958	24.555	66.28
199801:200012	96	-98.842	-100.839	-1.997	55.21
199901:200112	109	-85.088	-109.625	-24.537	75.23
200001:200212	119	-119.334	-260.565	-141.231	70.59
200101:200312	131	-446.355	-465.370	-19.015	64.12
200201:200412	133	-261.305	-285.984	-24.679	51.88
200301:200512	128	-105.933	-239.287	-133.353	49.22
200401:200612	131	-20.324	-169.740	-149.416	54.96
200501:200712	128	7.973	-116.063	-124.036	51.56
200601:200812	129	-129.902	-165.936	-36.034	78.29
200701:200912	129	19.864	-129.194	-149.058	34.11
200801:201012	126	-157.862	-179.444	-21.581	66.67
200901:201112	122	-259.939	-334.760	-74.821	55.74
201001:201212	121	-251.755	-234.360	17.395	67.77
201101:201312	121	-279.317	-205.105	74.211	66.94
201201:201412	119	-241.206	-273.286	-32.080	69.75
201301:201602	113	-320.076	-282.608	37.468	64.60
Average				-25.59	57.35

Panel D: Quartile 4(Carhart model)					
<i>Period</i>	<i># of Funds</i>	<i>Average</i> $\alpha_i^{*S\&P500}$ <i>(bp)</i>	<i>Average</i> α_i^{*Cat} <i>(bp)</i>	<i>Average difference</i> $\alpha_i^{*Cat} - \alpha_i^{*S\&P500}$ <i>(bp)</i>	<i>% Funds</i> <i>remaining in</i> <i>Quartile 4</i>
199201:199412	48	-680.588	-695.340	-14.752	72.92
199301:199512	55	-780.702	-735.892	44.809	69.09
199401:199612	61	-930.347	-884.578	45.769	67.21
199501:199712	69	-1436.729	-1173.304	263.425	69.57
199601:199812	75	-1073.939	-941.362	132.577	64.00
199701:199912	86	-1218.210	-1248.937	-30.727	80.23
199801:200012	96	-743.990	-740.403	3.588	68.75
199901:200112	108	-709.590	-800.757	-91.167	76.85
200001:200212	119	-839.593	-890.372	-50.779	68.91
200101:200312	132	-1038.163	-1096.031	-57.868	77.27
200201:200412	134	-770.944	-797.378	-26.434	79.10
200301:200512	129	-492.279	-635.970	-143.691	68.22
200401:200612	131	-445.152	-626.243	-181.091	78.63
200501:200712	129	-352.865	-477.177	-124.312	68.99
200601:200812	128	-517.257	-523.115	-5.858	80.47
200701:200912	128	-495.099	-480.083	15.015	53.91
200801:201012	127	-651.181	-620.318	30.863	81.10
200901:201112	123	-778.481	-777.586	0.895	65.04
201001:201212	122	-830.054	-835.978	-5.924	80.33
201101:201312	121	-1193.792	-802.262	391.529	69.42
201201:201412	118	-1137.554	-828.766	308.788	72.88
201301:201602	112	-832.393	-611.258	221.135	74.11
Average				32.99	72.14

Table 5: Percentage of funds that change performance quartiles: comparison of AGT with three-, four- and five – factors

The table shows percentage of funds that change performance quartiles when the benchmark changes from the S&P500 to the ‘category’ benchmark using the three- (FF3), four- (FF4) and five-(FF5) factors for the AGT model specification.

	FF3	FF4	FF5
Quartile 1:	32.40%	29.91%	26.53%
Quartile 2:	47.06%	44.50%	45.37%
Quartile 3:	45.42%	42.65%	44.26%
Quartile 4:	30.88%	28.86%	28.70%